On the Maximum Size of Block Codes Subject to a Distance Criterion

Vincent Y. F. Tan
National University of Singapore (NUS)



Ling-Hua Chang Yuan Ze Univ



Po-Ning Chen NCTU



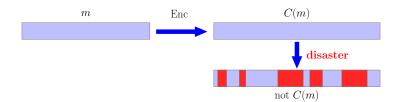
Carol Wang

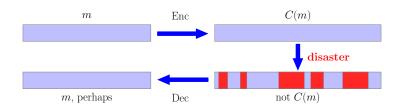


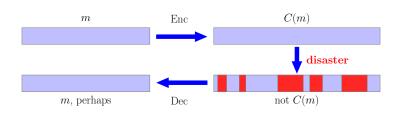
Yunghsiang Han Dongguan Univ. of Tech.

ITCom Workshop (Jan 2019)

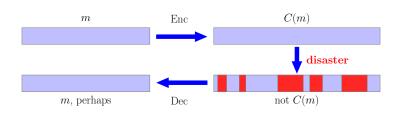




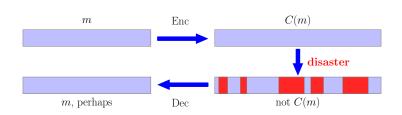




"Message" m (k symbols) maps to "codeword" C(m) (n > k symbols). Set of codewords is a code C.



"Message" m (k symbols) maps to "codeword" C(m) (n > k symbols). Set of codewords is a code C.

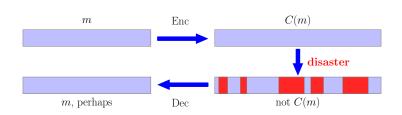


"Message" m (k symbols) maps to "codeword" C(m) (n > k symbols).

Set of codewords is a code C.

Key parameters:

■ Rate $\frac{1}{n} \log |\mathcal{C}|$: efficiency



"Message" m (k symbols) maps to "codeword" C(m) (n > k symbols).

Set of codewords is a code C.

Key parameters:

- Rate $\frac{1}{n} \log |\mathcal{C}|$: efficiency
- Distance : error-correction potential



Distance: "How many errors do we need to turn x into y?"

Distance: "How many errors do we need to turn x into y?"

Can correct as many errors as half the distance:

Distance: "How many errors do we need to turn x into y?"

Can correct as many errors as half the distance:

codeword	

```
codeword
```

Distance: "How many errors do we need to turn x into y?"

Can correct as many errors as half the distance:

codeword	
received	
codeword	



$$\mu(\mathbf{x}, \mathbf{y}) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1} \{ x_i \neq y_i \}$$
 (Hamming distance)

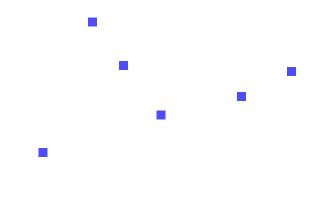
$$\mu(\mathbf{x}, \mathbf{y}) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1} \{ x_i \neq y_i \}$$
 (Hamming distance)
$$\mu(\mathbf{x}, \mathbf{y}) = \begin{cases} 0 & \mathbf{x} = \mathbf{y} \\ 1 & \text{else} \end{cases}$$
 (Probability-of-error distortion)

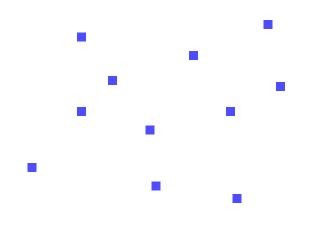
$$\mu(\mathbf{x},\mathbf{y}) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1} \{x_i \neq y_i\}$$
 (Hamming distance)
$$\mu(\mathbf{x},\mathbf{y}) = \begin{cases} 0 & \mathbf{x} = \mathbf{y} \\ 1 & \text{else} \end{cases}$$
 (Probability-of-error distortion)
$$\mu(\mathbf{x},\mathbf{y}) = \text{pretty much anything!}$$

$$\mu(\mathbf{x},\mathbf{y}) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1} \{x_i \neq y_i\}$$
 (Hamming distance)
$$\mu(\mathbf{x},\mathbf{y}) = \begin{cases} 0 & \mathbf{x} = \mathbf{y} \\ 1 & \text{else} \end{cases}$$
 (Probability-of-error distortion)
$$\mu(\mathbf{x},\mathbf{y}) = \text{pretty much anything!}$$
 (deletion distance, rank-metric, etc)



Question: What is the optimal rate-distance trade-off?





Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d = \delta n$ and rate $\approx 1 - H(\delta)$.

Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.

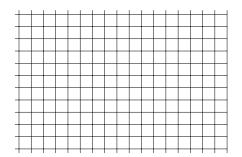
Proof 1: Greedy.

Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.

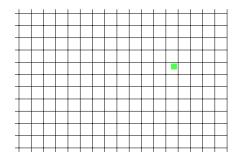
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



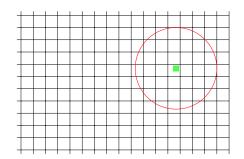
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



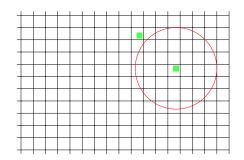
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



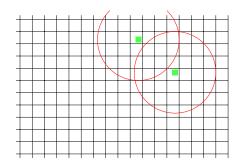
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d = \delta n$ and rate $\approx 1 - H(\delta)$.



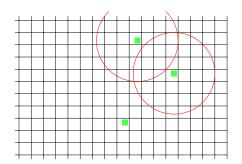
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



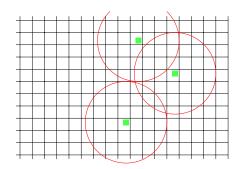
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



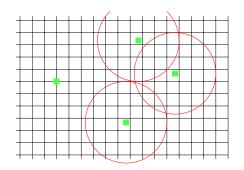
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



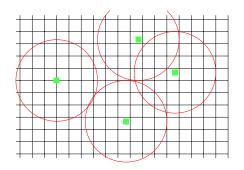
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



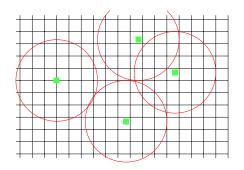
Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



Theorem (Gilbert-Varshamov bound)

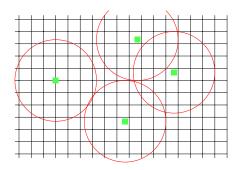
 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.



Theorem (Gilbert-Varshamov bound)

 \exists codes in $\{0,1\}^n$ with Hamming distance $d=\delta n$ and rate $\approx 1-H(\delta)$.

Proof 1: Greedy. Pick codewords at distance *d* until you can't.



Each circle has $\approx 2^{H(\delta)n}$ vectors, so final code size is $2^n/2^{H(\delta)n}$.

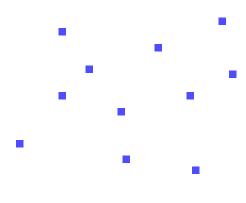
Proof 2: Random [Barg and Forney (2002)].

Proof 2: Random [Barg and Forney (2002)].

Pick i.i.d. codewords uniformly from $\{0,1\}^n$.

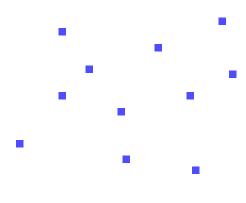
Proof 2: Random [Barg and Forney (2002)].

Pick i.i.d. codewords uniformly from $\{0,1\}^n$.



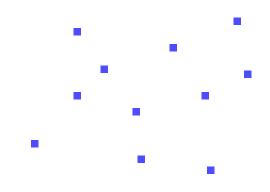
Proof 2: Random [Barg and Forney (2002)].

Pick i.i.d. codewords uniformly from $\{0,1\}^n$.



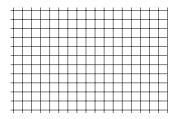
Proof 2: Random [Barg and Forney (2002)].

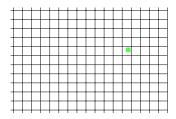
Pick i.i.d. codewords uniformly from $\{0,1\}^n$.

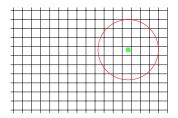


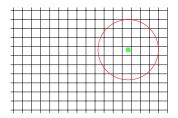
Works for rate $R \approx 1 - H(\delta)$ (proof on next slide).

Proof 2: Random.

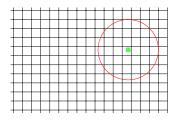






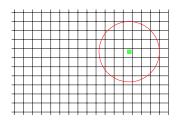


Proof 2: Random. Let $R = 1 - H(\delta) - \epsilon$.



Look at collision probability $\Pr[\mu(\mathbf{X}, \mathbf{Y}) < \delta n] = 2^{H(\delta)n}/2^n$.

Proof 2: Random. Let $R = 1 - H(\delta) - \epsilon$.

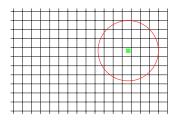


Look at collision probability $\Pr[\mu(\mathbf{X}, \mathbf{Y}) < \delta n] = 2^{H(\delta)n}/2^n$.

Number of "bad" pairs (x, y) is

$$\approx 2^{2Rn} \cdot \frac{2^{H(\delta)n}}{2^n} = 2^{(R-\epsilon)n}.$$

Proof 2: Random. Let $R = 1 - H(\delta) - \epsilon$.



Look at collision probability $\Pr[\mu(\mathbf{X}, \mathbf{Y}) < \delta n] = 2^{H(\delta)n}/2^n$.

Number of "bad" pairs (x, y) is

$$\approx 2^{2Rn} \cdot \frac{2^{H(\delta)n}}{2^n} = 2^{(R-\epsilon)n}.$$

Remove one element from each bad pair.

Distance is now δ , and rate is still $\approx R$.





Tightness of the GV bound is a major open question!

Tightness of the GV bound is a major open question!

This work: What if we don't use the *uniform* distribution in the random proof?

Tightness of the GV bound is a major open question!

This work: What if we don't use the *uniform* distribution in the random proof?

(Could imagine: supported on structured set, mixing distributions.)

Tightness of the GV bound is a major open question!

This work: What if we don't use the *uniform* distribution in the random proof?

(Could imagine: supported on structured set, mixing distributions.)

To mimic the GV proof, need to understand collision probability.

Tightness of the GV bound is a major open question!

This work: What if we don't use the *uniform* distribution in the random proof?

(Could imagine: supported on structured set, mixing distributions.)

To mimic the GV proof, need to understand collision probability.

When are two random codewords at distance < d?

Moral: For various **X**, want to understand collision probability (distance spectrum):

$$F_{\mathbf{X}}(d) := \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d],$$

where $\hat{\mathbf{X}}$ is an independent copy of \mathbf{X} .

Moral: For various **X**, want to understand collision probability (distance spectrum):

$$F_{\mathbf{X}}(d) := \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d],$$

where $\hat{\mathbf{X}}$ is an independent copy of \mathbf{X} .

Moral: For various **X**, want to understand collision probability (distance spectrum):

$$F_{\mathbf{X}}(d) := \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d],$$

where $\hat{\mathbf{X}}$ is an independent copy of \mathbf{X} .

$$F_{\mathbf{X}}(d) = \Pr[\mathbf{X} = \hat{\mathbf{X}}]$$

Moral: For various **X**, want to understand collision probability (distance spectrum):

$$F_{\mathbf{X}}(d) := \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d],$$

where $\hat{\mathbf{X}}$ is an independent copy of \mathbf{X} .

$$F_{\mathbf{X}}(d) = \Pr[\mathbf{X} = \hat{\mathbf{X}}]$$
$$= \sum_{\mathbf{x} \in \mathcal{C}} (P_{\mathbf{X}}(\mathbf{x}))^{2}$$

Moral: For various **X**, want to understand collision probability (distance spectrum):

$$F_{\mathbf{X}}(d) := \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d],$$

where $\hat{\mathbf{X}}$ is an independent copy of \mathbf{X} .

$$F_{\mathbf{X}}(d) = \Pr[\mathbf{X} = \hat{\mathbf{X}}]$$

$$= \sum_{\mathbf{x} \in \mathcal{C}} (P_{\mathbf{X}}(\mathbf{x}))^{2}$$

$$= \frac{1}{|\mathcal{C}|}.$$

Exact distance spectrum formula

Exact distance spectrum formula

So, if X is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

Exact distance spectrum formula

So, if X is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

So, if **X** is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

Theorem (Main theorem)

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right]}.$$

So, if X is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

Theorem (Main theorem)

Let $M^*(d)$ be the optimal size of a distance d code. Then

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

Key points:



So, if X is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

Theorem (Main theorem)

Let $M^*(d)$ be the optimal size of a distance d code. Then

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

Key points:

No asymptotics!

So, if **X** is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

Theorem (Main theorem)

Let $M^*(d)$ be the optimal size of a distance d code. Then

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

Key points:

- No asymptotics!
- Exact formula for basically any distance measure.



So, if **X** is uniform over C, then

$$|\mathcal{C}| = \frac{1}{F_{\mathbf{X}}(d)}.$$

In fact, this is tight.

Theorem (Main theorem)

Let $M^*(d)$ be the optimal size of a distance d code. Then

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

Key points:

- No asymptotics!
- Exact formula for basically any distance measure.
- Holds for arbitrary (non-discrete) alphabets.

Theorem

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

Theorem

Let $M^*(d)$ be the optimal size of a distance d code. Then

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

■ Turns question about codes into one about distributions.

Theorem

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

- Turns question about codes into one about distributions.
- Allows us to use optimization techniques for distributions.

Theorem

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right]}.$$

- Turns question about codes into one about distributions.
- Allows us to use optimization techniques for distributions.
- New bounds on the second-order asymptotics.

Theorem

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right]}.$$

- Turns question about codes into one about distributions.
- Allows us to use optimization techniques for distributions.
- New bounds on the second-order asymptotics.
- Best distribution is uniform over optimal code, but any distribution gives a lower bound.

For a fixed random vector **X**, want to show:

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \frac{1}{M^*(d)}.$$

For a fixed random vector **X**, want to show:

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \frac{1}{M^*(d)}.$$

Two steps:

1 If $|\operatorname{supp}(\mathbf{X})| = M \le M^*(d)$, then

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

For a fixed random vector **X**, want to show:

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \frac{1}{M^*(d)}.$$

Two steps:

If $|\operatorname{supp}(\mathbf{X})| = M \leq M^*(d)$, then

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

If $M > M^*(d)$, can reduce to first case.

We have

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2.$$

We have

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2.$$

Assume $|\operatorname{supp}(\mathbf{X})| = M \leq M^*(d)$.

We have

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2.$$

Assume $|\operatorname{supp}(\mathbf{X})| = M \leq M^*(d)$. Then

$$\frac{1}{M} \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x}) = \frac{1}{M}.$$

We have

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2.$$

Assume $|\operatorname{supp}(\mathbf{X})| = M \leq M^*(d)$. Then

$$\frac{1}{M} \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x}) = \frac{1}{M}.$$

By Cauchy-Schwartz,

$$\sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2 \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} \frac{1}{M^2} = \frac{1}{M} \ge \frac{1}{M^*(d)}.$$

We have

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2.$$

Assume $|\operatorname{supp}(\mathbf{X})| = M \le M^*(d)$. Then

$$\frac{1}{M} \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x}) = \frac{1}{M}.$$

By Cauchy-Schwartz,

$$\sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} P_{\mathbf{X}}(\mathbf{x})^2 \ge \sum_{\mathbf{x} \in \text{supp}(\mathbf{X})} \frac{1}{M^2} = \frac{1}{M} \ge \frac{1}{M^*(d)}.$$

So, for small support, uniform is best.





Showed that if $|\operatorname{supp}(\mathbf{X})|$ is small, $F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}$.

Showed that if $|\operatorname{supp}(\mathbf{X})|$ is small, $F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}$.

Idea: If $|\text{supp}(\mathbf{X})|$ is large, show how to reduce $|\text{supp}(\mathbf{X})|$ without increasing $F_{\mathbf{X}}(d)$.

Showed that if $|\operatorname{supp}(\mathbf{X})|$ is small, $F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}$.

Idea: If $|\text{supp}(\mathbf{X})|$ is large, show how to reduce $|\text{supp}(\mathbf{X})|$ without increasing $F_{\mathbf{X}}(d)$.

Specifically, we'll find X' with support size

$$|\operatorname{supp}(\mathbf{X})| - 1$$

and

$$F_{\mathbf{X}'}(d) \leq F_{\mathbf{X}}(d)$$
.

Showed that if $|\operatorname{supp}(\mathbf{X})|$ is small, $F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}$.

Idea: If $|\text{supp}(\mathbf{X})|$ is large, show how to reduce $|\text{supp}(\mathbf{X})|$ without increasing $F_{\mathbf{X}}(d)$.

Specifically, we'll find X' with support size

$$|\operatorname{supp}(\mathbf{X})| - 1$$

and

$$F_{\mathbf{X}'}(d) \leq F_{\mathbf{X}}(d).$$

If we iterate this until the support has size $M^*(d)$, then

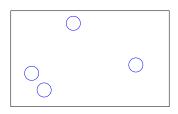
$$F_{\mathbf{X}}(d) \geq F_{\mathbf{X}'}(d) \geq F_{\mathbf{X}''}(d) \geq \cdots \geq \frac{1}{M^*(d)}.$$

Support reduction. Starting with distribution **X** on large support $M > M^*(d)$, want to construct **X**' on smaller support.

Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.

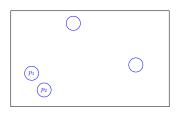
Support reduction. Starting with distribution **X** on large support $M > M^*(d)$, want to construct **X**' on smaller support.

Intuition
$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] = \sum_{i,j} p_i p_j \mathbf{1} \{\mu(\mathbf{x}_i, \mathbf{x}_j) < d\}$$
 where $p_i = P_{\mathbf{X}}(\mathbf{x}_i)$

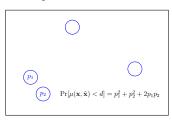


Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.

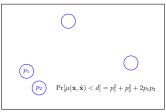
Intuition
$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] = \sum_{i,j} p_i p_j \mathbf{1} \{\mu(\mathbf{x}_i, \mathbf{x}_j) < d\}$$
 where $p_i = P_{\mathbf{X}}(\mathbf{x}_i)$

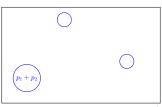


Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.

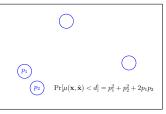


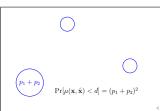
Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.





Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.





Support reduction. Starting with distribution **X** on large support $M > M^*(d)$, want to construct **X**' on smaller support.

Support reduction. Starting with distribution **X** on large support $M > M^*(d)$, want to construct **X**' on smaller support.

Proof.

If $|\text{supp}(\mathbf{X})| > M^*(d)$, have $\mathbf{x}, \mathbf{y} \in \text{supp}(\mathbf{X})$ at distance < d. Want to "combine" \mathbf{x}, \mathbf{y} .

Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.

Proof.

If $|\mathrm{supp}(\mathbf{X})| > M^*(d)$, have $\mathbf{x}, \mathbf{y} \in \mathrm{supp}(\mathbf{X})$ at distance < d. Want to "combine" \mathbf{x}, \mathbf{y} .

Question: Which of x, y to keep?

Support reduction. Starting with distribution **X** on large support $M > M^*(d)$, want to construct **X**' on smaller support.

Proof.

If $|\text{supp}(\mathbf{X})| > M^*(d)$, have $\mathbf{x}, \mathbf{y} \in \text{supp}(\mathbf{X})$ at distance < d. Want to "combine" \mathbf{x}, \mathbf{y} .

Question: Which of x, y to keep?

Answer: "Furthest": Keep x if

$$\Pr \big[\mu(\mathbf{x}, \mathbf{X}) < d \big] \leq \Pr \big[\mu(\mathbf{y}, \mathbf{X}) < d \big].$$

Support reduction. Starting with distribution X on large support $M > M^*(d)$, want to construct X' on smaller support.

Proof.

If $|\mathrm{supp}(\mathbf{X})| > M^*(d)$, have $\mathbf{x}, \mathbf{y} \in \mathrm{supp}(\mathbf{X})$ at distance < d. Want to "combine" \mathbf{x}, \mathbf{y} .

Question: Which of x, y to keep?

Answer: "Furthest": Keep x if

$$\Pr \big[\mu(\mathbf{x}, \mathbf{X}) < d \big] \le \Pr \big[\mu(\mathbf{y}, \mathbf{X}) < d \big].$$

Keeps distance spectrum (collision probability) $F_{\mathbf{X}}(d)$ small.



For **X** with small support,

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

For **X** with small support,

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

For other **X**, can reduce support size.

For **X** with small support,

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

For other X, can reduce support size.

Thus, optimal code size for distance *d* is

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

For **X** with small support,

$$F_{\mathbf{X}}(d) \geq \frac{1}{M^*(d)}.$$

For other **X**, can reduce support size.

Thus, optimal code size for distance d is

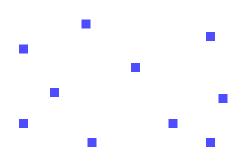
$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)} = \sup_{\mathbf{X}} \frac{1}{\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]}.$$

(Upper bound via uniform distribution.)

"Support reduction" proof is (sort of) constructive.

"Support reduction" proof is (sort of) constructive.

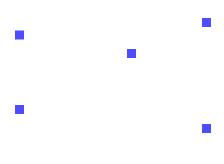
"Support reduction" proof is (sort of) constructive.



"Support reduction" proof is (sort of) constructive.



"Support reduction" proof is (sort of) constructive.

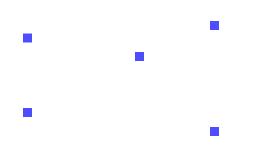


"Support reduction" proof is (sort of) constructive.



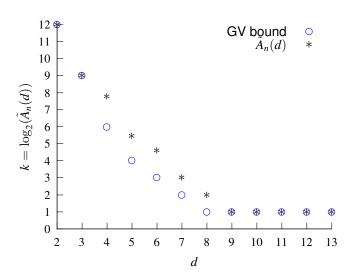
"Support reduction" proof is (sort of) constructive.

Start with any distribution, look at two codewords at distance < d, remove the one which is "closer" to the code.



Can be thought of as a different way to implement GV greedy construction. Seems to work well in simulations.

An Algorithmic Construction (n = 13)



■ Previous achievability proof only works for discrete (finite) alphabets because we used supp(X).

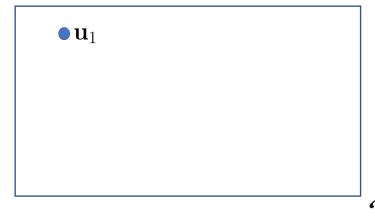
- Previous achievability proof only works for discrete (finite) alphabets because we used supp(X).
- Sort of similar to Motzkin-Strass (1965) and Korn (1968)
 - T. S. Motzkin and E. G. Straus, "Maxima for graphs and a new proof of a theorem of Turan," Canad. J. Math, vol. 17, no. 4, pp. 533–540, 1965.
 - 2 I. Korn, "On the lower bound of zero-error capacity," IEEE Trans. Inf. Theory, vol. 40, no. 4, pp. 509–510, May 1968.

- Previous achievability proof only works for discrete (finite) alphabets because we used supp(X).
- Sort of similar to Motzkin-Strass (1965) and Korn (1968)
 - T. S. Motzkin and E. G. Straus, "Maxima for graphs and a new proof of a theorem of Turan," Canad. J. Math, vol. 17, no. 4, pp. 533–540, 1965.
 - 2 I. Korn, "On the lower bound of zero-error capacity," IEEE Trans. Inf. Theory, vol. 40, no. 4, pp. 509–510, May 1968.
- We now generalize to the case in which $|\mathcal{X}| = \infty$ (even uncountable)

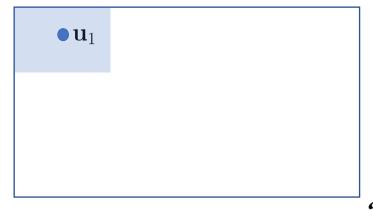
- Previous achievability proof only works for discrete (finite) alphabets because we used supp(X).
- Sort of similar to Motzkin-Strass (1965) and Korn (1968)
 - 1 T. S. Motzkin and E. G. Straus, "Maxima for graphs and a new proof of a theorem of Turan," Canad. J. Math, vol. 17, no. 4, pp. 533–540, 1965.
 - 2 I. Korn, "On the lower bound of zero-error capacity," IEEE Trans. Inf. Theory, vol. 40, no. 4, pp. 509–510, May 1968.
- We now generalize to the case in which $|\mathcal{X}| = \infty$ (even uncountable)
- Idea: Greedy selection of codewords $\{\mathbf{u}_i\}_{i=1}^k$ given a fixed random vector/distribution $\mathbf{X} \sim P_{\mathbf{X}}$.



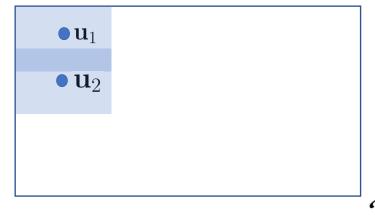




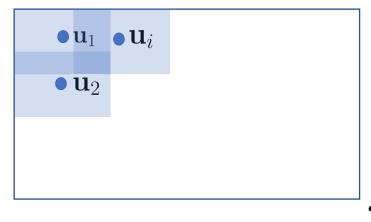
$$\mathbf{u}_1 = \operatorname{arg\,min}_{\mathbf{u}_1} \Pr \left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_1) \right]$$



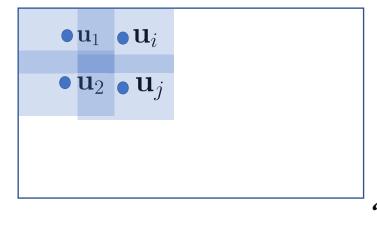
$$\mathbf{u}_1 = \operatorname{arg\,min}_{\mathbf{u}_1} \Pr \left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_1) \right]$$



$$\mathbf{u}_2 = \arg\min_{\mathbf{u}_2} \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_2) \setminus \mathcal{B}_d(\mathbf{u}_1)\right]$$

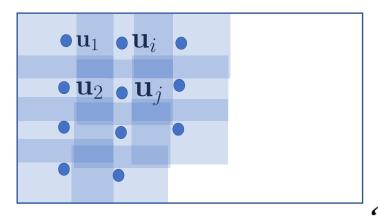


$$\mathbf{u}_i = \arg\min_{\mathbf{u}_i} \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \bigcup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right]$$



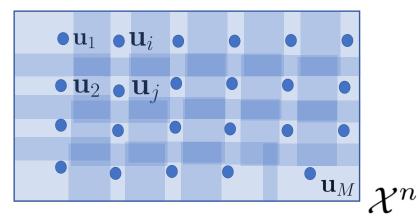
 \mathcal{X}^n

Choose more centers \mathbf{u}_{j} 's not in preceding balls.



 \mathcal{X}^n

And more balls...



Until you run out of space!

The code $C = \{\mathbf{u}_i : i = 1, \dots, M\}$ formed is a distance-d code and

$$p_j := \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right], \quad \text{satisfies} \quad \sum_{i=1}^M p_j = 1.$$

The code $C = \{\mathbf{u}_i : i = 1, \dots, M\}$ formed is a distance-d code and

$$p_j := \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right], \quad \text{satisfies} \quad \sum_{j=1}^M p_j = 1.$$

Let $\mathcal{D}_i := \mathcal{B}_d(\mathbf{u}_i) \setminus \bigcup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)$ and note that $\{\mathcal{D}_i\}$ forms a partition of \mathcal{X}^n .

The code $C = \{\mathbf{u}_i : i = 1, ..., M\}$ formed is a distance-d code and

$$p_j := \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right], \quad \text{satisfies} \quad \sum_{j=1}^M p_j = 1.$$

Let $\mathcal{D}_i := \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)$ and note that $\{\mathcal{D}_i\}$ forms a partition of \mathcal{X}^n .

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] = \sum_{j=1}^{M} \int_{\mathbf{x} \in \mathcal{D}_{j}} \left(\int_{\hat{\mathbf{x}} \in \mathcal{B}_{d}(\mathbf{x})} dP_{\mathbf{X}}(\hat{\mathbf{x}}) \right) dP_{\mathbf{X}}(\mathbf{x}) \quad :: \mathbf{X} \perp \!\!\! \perp \hat{\mathbf{X}}$$

The code $C = \{\mathbf{u}_i : i = 1, ..., M\}$ formed is a distance-d code and

$$p_j := \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right], \quad \text{satisfies} \quad \sum_{j=1}^M p_j = 1.$$

Let $\mathcal{D}_i := \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)$ and note that $\{\mathcal{D}_i\}$ forms a partition of \mathcal{X}^n .

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] = \sum_{j=1}^{M} \int_{\mathbf{x} \in \mathcal{D}_{j}} \left(\int_{\hat{\mathbf{x}} \in \mathcal{B}_{d}(\mathbf{x})} dP_{\mathbf{X}}(\hat{\mathbf{x}}) \right) dP_{\mathbf{X}}(\mathbf{x}) \quad \therefore \quad \mathbf{X} \perp \perp \hat{\mathbf{X}}$$

$$\geq \sum_{j=1}^{M} \int_{\mathbf{x} \in \mathcal{D}_{j}} p_{j} dP_{\mathbf{X}}(\mathbf{x}) \quad \therefore \min_{\mathbf{x} \in \mathcal{D}_{j}} P_{\mathbf{X}} \{\mathcal{B}_{d}(\mathbf{x})\} \geq p_{j}$$

Non-Discrete Code Alphabets: Achievability Proof

The code $C = \{\mathbf{u}_i : i = 1, ..., M\}$ formed is a distance-d code and

$$p_j := \Pr\left[\mathbf{X} \in \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)\right], \quad \text{satisfies} \quad \sum_{j=1}^M p_j = 1.$$

Let $\mathcal{D}_i := \mathcal{B}_d(\mathbf{u}_i) \setminus \cup_{j=1}^{i-1} \mathcal{B}_d(\mathbf{u}_j)$ and note that $\{\mathcal{D}_i\}$ forms a partition of \mathcal{X}^n .

$$\begin{split} \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d] &= \sum_{j=1}^{M} \int_{\mathbf{x} \in \mathcal{D}_{j}} \left(\int_{\hat{\mathbf{x}} \in \mathcal{B}_{d}(\mathbf{x})} dP_{\mathbf{X}}(\hat{\mathbf{x}}) \right) dP_{\mathbf{X}}(\mathbf{x}) \quad \because \mathbf{X} \perp \!\!\! \perp \hat{\mathbf{X}} \\ &\geq \sum_{j=1}^{M} \int_{\mathbf{x} \in \mathcal{D}_{j}} p_{j} dP_{\mathbf{X}}(\mathbf{x}) \quad \because \min_{\mathbf{x} \in \mathcal{D}_{j}} P_{\mathbf{X}} \{\mathcal{B}_{d}(\mathbf{x})\} \geq p_{j} \\ &\geq \sum_{j=1}^{M} p_{j}^{2} \geq \frac{1}{M} \geq \frac{1}{M^{*}(d)} \quad \because \text{ Cauchy-Schwarz & } M \leq M^{*}(d) \end{split}$$

 Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)

- Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)
- But we removed space $\mathcal{B}_d(\mathbf{u}_k) \subset \mathcal{X}^n$ successively instead of codewords successively.

- Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)
- But we removed space $\mathcal{B}_d(\mathbf{u}_k) \subset \mathcal{X}^n$ successively instead of codewords successively.
- Showed through simple algebraic manipulations that for any X,

$$F_{\mathbf{X}}(d) = \Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right] \ge \frac{1}{M^*(d)}$$

- Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)
- But we removed space $\mathcal{B}_d(\mathbf{u}_k) \subset \mathcal{X}^n$ successively instead of codewords successively.
- Showed through simple algebraic manipulations that for any X,

$$F_{\mathbf{X}}(d) = \Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right] \ge \frac{1}{M^*(d)} \quad \Longrightarrow \quad M^*(d) \ge \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)}.$$

- Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)
- But we removed space $\mathcal{B}_d(\mathbf{u}_k) \subset \mathcal{X}^n$ successively instead of codewords successively.
- Showed through simple algebraic manipulations that for any X,

$$F_{\mathbf{X}}(d) = \Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right] \ge \frac{1}{M^*(d)} \implies M^*(d) \ge \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)}.$$

■ Converse part is the same as for discrete alphabets (hinges on uniform distribution over optimal code \mathcal{C}^*)

- Also used a greedy construction (à la Feinstein's lemma in information spectrum analysis)
- But we removed space $\mathcal{B}_d(\mathbf{u}_k) \subset \mathcal{X}^n$ successively instead of codewords successively.
- Showed through simple algebraic manipulations that for any X,

$$F_{\mathbf{X}}(d) = \Pr\left[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d\right] \ge \frac{1}{M^*(d)} \implies M^*(d) \ge \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)}.$$

- Converse part is the same as for discrete alphabets (hinges on uniform distribution over optimal code \mathcal{C}^*)
- In summary,

$$M^*(d) = \sup_{\mathbf{X}} \frac{1}{F_{\mathbf{X}}(d)}$$

Corollary (Refined GV bound)

For the Hamming distance, the optimal code rate for distance δn is

$$R_n^*(\delta) \ge 1 - H(\delta) + \frac{\log n}{2n} + \Theta\left(\frac{1}{n}\right).$$

Corollary (Refined GV bound)

For the Hamming distance, the optimal code rate for distance δn is

$$R_n^*(\delta) \ge 1 - H(\delta) + \frac{\log n}{2n} + \Theta\left(\frac{1}{n}\right).$$

Proof.

Let **X** be uniform on $\{0,1\}^n$.

Corollary (Refined GV bound)

For the Hamming distance, the optimal code rate for distance δn is

$$R_n^*(\delta) \ge 1 - H(\delta) + \frac{\log n}{2n} + \Theta\left(\frac{1}{n}\right).$$

Proof.

Let **X** be uniform on $\{0,1\}^n$.

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < \delta n] = \Pr\left[\frac{1}{n} \sum_{i=1}^{n} \mathbf{1}\{X_i \neq \hat{X}_i\} < \delta\right] \sim c \cdot \frac{2^{n[1-H(\delta)]}}{\sqrt{n}}.$$

Result follows using exact asymptotics for sums of i.i.d. variables.



Corollary (Refined GV bound)

For the Hamming distance, the optimal code rate for distance δn is

$$R_n^*(\delta) \ge 1 - H(\delta) + \frac{\log n}{2n} + \Theta\left(\frac{1}{n}\right).$$

Proof.

Let **X** be uniform on $\{0,1\}^n$.

$$\Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < \delta n] = \Pr\left[\frac{1}{n} \sum_{i=1}^{n} \mathbf{1}\{X_i \neq \hat{X}_i\} < \delta\right] \sim c \cdot \frac{2^{n[1-H(\delta)]}}{\sqrt{n}}.$$

Result follows using exact asymptotics for sums of i.i.d. variables.

Jiang and Vardy (2004) showed that the "second-order term" $\geq \frac{\log n}{n}$.

Corollary (Upper Bound on Rate)

For any arbitrary bounded distance measure, the optimal code rate for distance δn is

$$R_n^*(\delta) \leq I_{X^n}(\delta) + O\left(\frac{1}{\sqrt{n}}\right).$$

where the large-deviations rate function is

$$I_{X^n}(a) := \sup_{ heta} \left\{ a heta - arphi_{X^n}(heta)
ight\}, \quad extit{and} \quad arphi_X(heta) := \log \mathbb{E} \left[e^{ heta \mu(X,\hat{X})}
ight].$$

Corollary (Upper Bound on Rate)

For any arbitrary bounded distance measure, the optimal code rate for distance δn is

$$R_n^*(\delta) \leq I_{X^n}(\delta) + O\left(\frac{1}{\sqrt{n}}\right).$$

where the large-deviations rate function is

$$I_{X^n}(a) := \sup_{ heta} \left\{ a heta - arphi_{X^n}(heta)
ight\}, \quad extit{and} \quad arphi_X(heta) := \log \mathbb{E} \left[e^{ heta \mu(X,\hat{X})}
ight].$$

Proof.

Careful tilting of probability distributions.



First-Order Asymptotics

First-Order Asymptotics

Corollary (First-Order Asymptotics on Rate)

If the sequence of distance measures satisfies

$$\sup_{n\in\mathbb{N}}\max_{x^n,\hat{x}^n}\frac{1}{n}\mu(x^n,\hat{x}^n)<\infty,$$

then we have

$$\limsup_{n \to \infty} R_n^*(\delta) = \limsup_{n \to \infty} I_{X^n}(\delta), \quad \textit{and}$$
 $\liminf_{n \to \infty} R_n^*(\delta) = \liminf_{n \to \infty} I_{X^n}(\delta)$

where the large-deviations rate function is

$$I_{X^n}(a) := \sup_{ heta} \left\{ a heta - arphi_{X^n}(heta)
ight\}, \quad ext{and} \quad arphi_X(heta) := \log \mathbb{E} \left[e^{ heta \mu(X,\hat{X})}
ight].$$

Corollary (Hamming Bound for Finite $|\mathcal{X}|$)

$$M^*(d) \leq \inf_{\epsilon > 0} \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{(d-\epsilon)/2}(\mathbf{0})\right|} \leq \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{\lfloor (d-1)/2 \rfloor}(\mathbf{0})\right|}$$

Corollary (Hamming Bound for Finite $|\mathcal{X}|$)

$$M^*(d) \leq \inf_{\epsilon>0} \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{(d-\epsilon)/2}(\mathbf{0})\right|} \leq \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{\lfloor (d-1)/2 \rfloor}(\mathbf{0})\right|}$$

Proof: (Due to V. Guruswami).

Let
$$e = (d - \epsilon)/2$$
. Then

$$|\mathcal{B}_e(\mathbf{0})| F_{\mathbf{X}}(d) = \sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_e(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) \sum_{\mathbf{z}: \mu(\mathbf{x}, \mathbf{z}) < d} P_{\mathbf{X}}(\mathbf{z})$$

Corollary (Hamming Bound for Finite $|\mathcal{X}|$)

$$M^*(d) \le \inf_{\epsilon > 0} \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{(d-\epsilon)/2}(\mathbf{0})\right|} \le \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{\lfloor (d-1)/2 \rfloor}(\mathbf{0})\right|}$$

Proof: (Due to V. Guruswami).

Let
$$e = (d - \epsilon)/2$$
. Then

$$\begin{aligned} |\mathcal{B}_{e}(\mathbf{0})|F_{\mathbf{X}}(d) &= \sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_{e}(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) \sum_{\mathbf{z}: \mu(\mathbf{x}, \mathbf{z}) < d} P_{\mathbf{X}}(\mathbf{z}) \\ &\geq \sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_{e}(\mathbf{x})} \sum_{\mathbf{z} \in \mathcal{B}_{e}(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) P_{\mathbf{X}}(\mathbf{z}) \end{aligned}$$

Corollary (Hamming Bound for Finite $|\mathcal{X}|$)

$$M^*(d) \leq \inf_{\epsilon>0} \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{(d-\epsilon)/2}(\mathbf{0})\right|} \leq \frac{|\mathcal{X}|^n}{\left|\mathcal{B}_{\lfloor (d-1)/2 \rfloor}(\mathbf{0})\right|}$$

Proof: (Due to V. Guruswami).

Let
$$e = (d - \epsilon)/2$$
. Then

$$\begin{split} |\mathcal{B}_{e}(\mathbf{0})|F_{\mathbf{X}}(d) &= \sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_{e}(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) \sum_{\mathbf{z}: \mu(\mathbf{x}, \mathbf{z}) < d} P_{\mathbf{X}}(\mathbf{z}) \\ &\geq \sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_{e}(\mathbf{x})} \sum_{\mathbf{z} \in \mathcal{B}_{e}(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) P_{\mathbf{X}}(\mathbf{z}) \\ &\stackrel{\text{CS}}{\geq} \left(\sum_{\mathbf{x}} \sum_{\mathbf{y} \in \mathcal{B}_{e}(\mathbf{x})} P_{\mathbf{X}}(\mathbf{y}) \right)^{2} = \frac{|\mathcal{B}_{e}(\mathbf{0})|^{2}}{|\mathcal{X}|^{n}} \end{split}$$



IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

809

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluates.

surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)/n}$.

IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

007

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minimum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluated

surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)}/n$.



My visit to NCTU 2015

IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

809

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minimum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluated

surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)}/n$.



My visit to NCTU 2015

Chen, Lee and Han (2000) proved an elegant information spectrum-style result

IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

809

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minimum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluated surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)}/n$.



My visit to NCTU 2015

Chen, Lee and Han (2000) proved an elegant information spectrum-style result

$$\limsup_{n\to\infty} \delta_n^*(2^{nR}) = \sup_{\mathbf{X}=\{X^n\}_{n=1}^\infty} \overline{\Lambda}_{\mathbf{X}}(R) \quad \text{(except at countably many points)}$$

IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

869

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minmum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluated surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)}/n$.



My visit to NCTU 2015

■ Chen, Lee and Han (2000) proved an elegant information spectrum-style result

$$\limsup_{n\to\infty} \delta_n^*(2^{nR}) = \sup_{\mathbf{X} = \{X^n\}_{n=1}^\infty} \overline{\Lambda}_{\mathbf{X}}(R) \quad \text{(except at countably many points)}$$

$$\overline{\Lambda}_{\mathbf{X}}(R) := \inf\Big\{a \in \mathbb{R} : \lim_{n\to\infty} \Pr\big[\mu(X^n, \hat{X}^n) > a\big]^{2^{nR}} = 0\Big\}.$$

IEEE TRANSACTIONS ON INFORMATION THEORY, VOL. 46, NO. 3, MAY 2000

Distance-Spectrum Formulas on the Largest Minimum Distance of Block Codes

Po-Ning Chen, Member, IEEE, Tzong-Yow Lee, and Yunghsiang S. Han, Member, IEEE

Abstract—A general formula for the asymptotic largest minimum distance (in block length) of deterministic block codes under generalized distance functions (not necessarily additive, symmetric, and bounded) is presented. As revealed in the formula, the largest minimum distance can be fully determined by the ultimate statistical characteristics of the normalized distance function evaluated surable function on the "distance" between two code symbols, determine the asymptotic ratio, the largest minimum distance attainable among M selected codewords divided by the code block length n, as n tends to infinity, subject to a fixed rate $R \triangleq \log{(M)}/n$.



My visit to NCTU 2015

■ Chen, Lee and Han (2000) proved an elegant information spectrum-style result

$$\limsup_{n\to\infty} \delta_n^*(2^{nR}) = \sup_{\mathbf{X} = \{X^n\}_{n=1}^\infty} \overline{\Lambda}_{\mathbf{X}}(R) \quad \text{(except at countably many points)}$$

$$\overline{\Lambda}_{\mathbf{X}}(R) := \inf \Big\{ a \in \mathbb{R} : \lim_{n \to \infty} \Pr \left[\mu(X^n, \hat{X}^n) > a \right]^{2^{nR}} = 0 \Big\}.$$

■ The present result is a non-asymptotic version of CLH2000.



 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

Some open questions.

Better algorithm (improved rule for combining codewords)?

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

Some open questions.

- Better algorithm (improved rule for combining codewords)?
- Better bounds for the current algorithm?

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

Some open questions.

- Better algorithm (improved rule for combining codewords)?
- Better bounds for the current algorithm?
- Improved codes?

 Showed how to connect optimal code size/distance tradeoff and distance spectrum

$$F_{\mathbf{X}}(d) = \Pr[\mu(\mathbf{X}, \hat{\mathbf{X}}) < d]$$

for different random vectors X.

Also got an algorithm for constructing codes.

Some open questions.

- Better algorithm (improved rule for combining codewords)?
- Better bounds for the current algorithm?
- Improved codes?
- To appear in the IEEE Transactions on Information Theory in 2019.



Thanks!



Thanks!



My collaborators and I at ITW 2017 (Kaohsiung)

