Recent Advances in Ranking: Adversarial Respondents and Lower Bounds on the Bayes Risk

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Ranking

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1 Introduction to Statistical Models for Ranking

2 Fundamental Limits of Top-K Ranking with Adversaries

3 Lower Bounds on the Bayes Risk of a Bayesian BTL Model

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A fundamental problem in a wide range of contexts

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- Applications: web search, recommendation systems, social choice, sports competitions, voting, etc.
- Efforts in developing various ranking algorithms
- A variety of statistical models introduced for evaluating ranking algorithms

Ranking: An Example and Difficulties

Example: Web search

Google YAHOO! bing

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Ranking: An Example and Difficulties

Example: Web search

Google YAHOO! bing



- $n = 10^9$ websites
- $\binom{n}{2} \approx n^2 = 10^{18}$ comparisons

• Do we really need $\Theta(n^2)$ comparisons?

Suppose that

- we want a total ordering
- pairwise comparisons are randomly given (probabilistically).

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- No way to identify the ordering between 1 and 2 without a direct comparison, i.e., comparison must be made w.p. 1
- Worse with noisy data
- Adopt a Shannon-theoretic approach in our analyses



Top-K Ranking Usually Suffices

Huge number of movies



Find only top K = 3



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Top-*K* Ranking Usually Suffices

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Find only top K = 3





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Adopt the Bradley-Terry-Luce or BTL model in which there is an underlying unknown score vector

$$\mathbf{w} = (w_1, \ldots, w_n) \in \mathbb{R}^n_{++},$$

where w_i is the likeability of movie *i*.

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Decide which items to compare via a comparison graph





■ The outcome of the comparison between item 1 and 2 is

$$Y_{12} = \mathbb{I}\{\text{item 1} \succ \text{item 2}\} \sim \text{Bern}\left(\frac{w_1}{w_1 + w_2}\right).$$



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■ E.g., $w_1 = 0.9$ and $w_2 = 0.1$, then item 1 beats item 2 w.p. 90%.



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We have L independent copies

$$Y_{ij}^{(1)},\ldots,Y_{ij}^{(L)}$$

for each observed edge $\{i, j\} \in \mathcal{E}$ of the observation graph.



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■ We have *L* independent copies

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Determine fundamental limits on L (as a function of n and other parameters) so that recovery of top-K set is successful.

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Top-*K* Ranking with Adversaries

Joint work with





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Top-K Ranking with Adversaries

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■ C. Suh, **VYFT** and R. Zhao "Adversarial Top-*K* Ranking", IEEE Trans. on Inf. Theory, Apr 2017

faithful population η portion



adversarial population

 $1-\eta$ portion



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 $\begin{array}{c} {\rm faithful} \\ {\rm population} \\ \eta \\ {\rm portion} \end{array}$



$$Y_{ij} \sim \operatorname{Bern}\left(\frac{w_i}{w_i + w_j}\right)$$

adversarial population

 $1-\eta$ portion

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Spammers provide answers in an adversarial manner



Spammers provide answers in an adversarial manner

$$Y_{ij} \sim \text{Bern}\left(\eta \cdot \frac{w_i}{w_i + w_j} + (1 - \eta) \cdot \frac{w_j}{w_i + w_j}\right)$$

Given an observed pair, each sample has different distributions

$$Y_{ij}^{(l)} \sim \text{Bern}\left(\frac{\eta_l}{w_i + w_j} + (1 - \eta_l) \cdot \frac{1/w_i}{1/w_i + 1/w_j}\right)$$

where η_l is a quality parameter of measurement l

¹X. Chen, P. N. Bennett, K. Collins-Thompson, and E. Horvitz, "Pairwise ranking aggregation in a crowdsourced setting," in WSDM, 2013

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- Subsumes as a special case our adversarial BTL model when all quality parameters are the same
- The authors developed a ranking algorithm but without theoretical guarantees
- More difficult to analyze as there are many more parameters

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Goal of Adversarial Top-K Ranking

Erdös-Rényi comparison graph



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Erdös-Rényi comparison graph




η = Fraction of non-adversaries; $\Delta_K \approx w_K - w_{K+1}$

sample complexity



 $\eta = 1$ studied by Chen and Suh (2015)²

²Y. Chen and C. Suh, "Spectral MLE: Top-K rank aggregation from pairwise comparisons," in ICML 2015

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Experimental Results for n = 1000 and K = 10



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Optimality



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Image: A matrix



Minimax optimality: Construct "worst-case" score vectors



- Minimax optimality: Construct "worst-case" score vectors
- Translation to *M*-ary hypothesis testing: Construction of multiple hypotheses



- Minimax optimality: Construct "worst-case" score vectors
- Translation to *M*-ary hypothesis testing: Construction of multiple hypotheses
- Information-theoretic ideas applied to statistical learning

Construction of $M := \min\{K, n - K\} + 1 \le n/2$ hypotheses:

$$\Pr(\sigma([K]) = S) = \frac{1}{M}, \text{ for } S = \{2, \dots, K\} \cup \{i\}, i = 1, K + 1, \dots, n\}$$

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Bound mutual info. of permutation and "erased" version of $Y_{ii}^{(l)}$:

$$I(\sigma; \mathbf{Z}) \leq \frac{p}{M^2} \sum_{\sigma_1, \sigma_2 \in \mathcal{M}} \sum_{l=1}^{L} \left\{ \sum_{i \neq j} D\left(P_{Y_{ij}^{(l)} | \sigma_1} \left\| P_{Y_{ij}^{(l)} | \sigma_2} \right) \right\}$$

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Bound the divergence using reverse Pinsker's inequality. Here is where Δ_K comes in

$$\sum_{i \neq j} D\left(P_{Y_{ij}^{(l)}|\sigma_1} \left\| P_{Y_{ij}^{(l)}|\sigma_2}\right) \le n \cdot (2\eta - 1)^2 \cdot \Delta_K^2$$

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Fano's inequality

Ranking Algorithm for η Known: Part I



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Ranking Algorithm for η Known: Part I



Scores determine the ranking

Ranking Algorithm for η Known: Part I



- Scores determine the ranking
- Adopt a two-step approach

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Ranking Algorithm for η Known: Part II



Ranking Algorithm for η Known: Part II



Key Message:

Small MSE \Longrightarrow Small ℓ_{∞} Error of $\hat{w} \implies$ High Ranking Accuracy



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- Recall $\eta = 1$ (no adversaries)
- *L* independent copies $Y_{ij}^{(1)}, Y_{ij}^{(2)}, \ldots, Y_{ij}^{(L)}$



- **Recall** $\eta = 1$ (no adversaries)
- L independent copies $Y_{ij}^{(1)}, Y_{ij}^{(2)}, \dots, Y_{ij}^{(L)}$
- Convergence to stationary distribution

$$\frac{1}{L}\sum_{l=1}^{L}Y_{ij}^{(l)} \to \frac{w_i}{w_i + w_j}$$



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Detailed balance equation:

$$\pi_i \cdot \frac{w_j}{w_i + w_j} = \pi_j \cdot \frac{w_i}{w_i + w_j}$$

where $\boldsymbol{\pi} := [\pi_1, \pi_2, \dots, \pi_n]$ is the stat. distn. of the chain.



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Stationary distribution converges to w (up to constant scaling), i.e.,

$$\lim_{L\to\infty}\boldsymbol{\pi}^{(L)} = \alpha \mathbf{w}$$



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• Arbitrary $\eta \in (1/2, 1]$ (adversaries)

- *L* independent copies $Y_{ij}^{(1)}, Y_{ij}^{(2)}, \ldots, Y_{ij}^{(L)}$
- Redefine Markov chain



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$$\eta \in (1/2, 1]$$
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- Redefine Markov chain

We instead have the following convergence:

$$\frac{1}{L}\sum_{l=1}^{L}Y_{ij}^{(l)} \to \eta \frac{w_i}{w_i + w_j} + (1 - \eta)\frac{w_j}{w_i + w_j} = (2\eta - 1)\frac{w_i}{w_i + w_j} + (1 - \eta)$$



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Redefine "shifted" samples with range scaled by $2\eta - 1$:

$$\tilde{Y}_{ij} = \frac{1}{2\eta - 1} \left[\frac{1}{L} \sum_{l=1}^{L} Y_{ij}^{(l)} - (1 - \eta) \right] \rightarrow \frac{w_i}{w_i + w_j}$$



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Construct Markov chain with transition probabilities $\{\tilde{Y}_{ij}\}$.

Ranking Algorithm for η Known: Summary



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Ranking Algorithm for η Known: Summary



Use several concentraition inequalities (Hoeffding, Bernstein, Tropp, etc.), we can show that if

sample size
$$= L\binom{n}{2}p \succeq \frac{n\log n}{(2\eta - 1)^2 \Delta_K^2} \implies$$
 Feasible Top-*K* Ranking

What if η is unknown?

Adversarial BTL model is a mixture model

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³P. Jain and S. Oh, "Learning mixtures of discrete product distributions using spectral decompositions," in COLT, 2014

⁴A. Anandkumar, R. Ge, D. Hsu, S. M. Kakade, and M. Telgarsky, "Tensor decompositions for learning latent variable models," JMLR, 2014 A Control of the second se
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- Obtaining global optimality guarantees for mixture model problems is difficult in general

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- Recent developments:
 - Tensor methods: Jain and Oh³ and Anandkumar et al.⁴
 - Key idea: Exact 2nd and 3rd moments yield sufficient statistics

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- Recent developments:
 - Tensor methods: Jain and Oh³ and Anandkumar et al.⁴
 - Key idea: Exact 2nd and 3rd moments yield sufficient statistics
- Our setting:
 - Can obtain estimates of 2nd and 3rd moments
 - Can estimate η

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1 Turn weights into distribution vectors

$$\pi_0 = \begin{bmatrix} \cdots & \frac{w_i}{w_i + w_j} & \frac{w_j}{w_i + w_j} & \frac{w_{i'}}{w_{i'} + w_{j'}} & \frac{w_{j'}}{w_{i'} + w_{j'}} & \cdots \end{bmatrix}^T$$

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2 Estimate moments. Ground truth moment matrix and tensor are:

$$\begin{split} M_2 &:= \eta \pi_0 \otimes \pi_0 + (1 - \eta) \pi_1 \otimes \pi_1, \\ M_3 &:= \eta \pi_0 \otimes \pi_0 \otimes \pi_0 + (1 - \eta) \pi_1 \otimes \pi_1 \otimes \pi_1. \end{split}$$

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3 Solves a Least Squares Problem

$$\hat{G} \in \operatorname*{arg\,min}_{Z \in \mathbb{R}^{2 \times 2 \times 2}} \left\| \mathcal{P}_{\Omega_3} \left(Z \left[P_{\hat{M}_2} \right]_3 - \frac{1}{|\mathcal{I}_2|} \sum_{t \in \mathcal{I}_2} \otimes^3 \underline{Y}^{(t)} \right) \right\|_F^2$$

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4 Find leading eigenvalue $\lambda_1(\hat{G})$ of \hat{G} which is related to η as follows:

$$\hat{\eta} = \lambda_1 (\hat{G})^{-2}$$

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How does the quality of the estimation of η affect overall sample complexity?

With very careful analysis, we can derive a meta-lemma

$$|\hat{\eta} - \eta| \le \epsilon \implies \text{Sample size} = L\binom{n}{2}p \succeq \frac{n\log^2 n}{\epsilon^2}$$

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This implies that

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 implies that $\|\hat{\mathbf{w}} - \mathbf{w}\|_{\infty} \downarrow$ but sample size \uparrow

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 $\| \hat{\eta} - \eta \| \downarrow$ implies that $\| \hat{\mathbf{w}} - \mathbf{w} \|_{\infty} \downarrow$ but sample size \uparrow

■
$$|\hat{\eta} - \eta|$$
 ↑ implies that sample size ↓ but $\|\hat{\mathbf{w}} - \mathbf{w}\|_{\infty}$ ↑

With very careful analysis, we can derive a meta-lemma

$$|\hat{\eta} - \eta| \le \epsilon \implies \text{Sample size} = L\binom{n}{2}p \succeq \frac{n\log^2 n}{\epsilon^2}$$

- This implies that
 - $\| \hat{\eta} \eta \| \downarrow \text{ implies that } \| \hat{\mathbf{w}} \mathbf{w} \|_{\infty} \downarrow \text{ but sample size } \uparrow$
 - $\blacksquare |\hat{\eta} \eta| \uparrow \text{ implies that sample size } \downarrow \text{ but } \|\hat{\mathbf{w}} \mathbf{w}\|_{\infty} \uparrow$
- Find a sweet spot to show that

sample size
$$\succeq \frac{n \log^2 n}{(2\eta - 1)^4 \Delta_K^4}$$
, \implies Feasible Top-*K* ranking

Explored a Top-K ranking problem for an adversarial setting

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- Explored a Top-K ranking problem for an adversarial setting
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- Explored a Top-K ranking problem for an adversarial setting
- Characterized exact order-wise optimal sample complexity for η -known case
- Established an upper bound on the sample complexity for the η -unknown case
- Developed computationally efficient algorithms for both cases (using state-of-the-art tensor methods for the η-unknown case)

- Explored a Top-K ranking problem for an adversarial setting
- Characterized exact order-wise optimal sample complexity for η -known case
- Established an upper bound on the sample complexity for the η -unknown case
- Developed computationally efficient algorithms for both cases (using state-of-the-art tensor methods for the η-unknown case)
- C. Suh, **VYFT** and R. Zhao "Adversarial Top-*K* Ranking", IEEE Trans. on Inf. Theory, Apr 2017

1 Introduction to Statistical Models for Ranking

2 Fundamental Limits of Top-*K* Ranking with Adversaries

3 Lower Bounds on the Bayes Risk of a Bayesian BTL Model

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Lower Bounds on the Risk of a Bayesian BTL Model

Joint work with





Mine Alsan Ranjitha Prasad (NUS) (TCS Innovation Labs, Delhi)

Lower Bounds on the Risk of a Bayesian BTL Model

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M. Alsan, R. Prasad and VYFT, "Lower Bounds on the Bayes Risk of the Bayesian BTL Model with Applications to Comparison Graphs", IEEE J. on Sel. Topics of Sig. Proc., Oct 2018

Summary of contributions

Study the fundamental performance limits of ranking algorithms in the Bradley-Terry-Luce model within a Bayesian framework:

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for norm-based distortion functions $\|\cdot\|_r^r$, for any $r \ge 1$.

- The Bayesian Cramér-Rao bound for the MSE, i.e., r = 2.

2 Explore optimal comparison graph structures to design experiments minimizing distortion.

BTL model: To each item $i \in [n]$, a skill parameter $w_i \in \mathbb{R}_{++}$ s.t.

$$P_{ij} := \Pr[\operatorname{item} i \succ \operatorname{item} j] = \frac{w_i}{w_i + w_j}.$$

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indep. pairwise comparisons, we count:

m_{ij}: Num. of pairwise comparisons between items *i* & *j*,
b_{ij}: Num. of comparisons in which *i* is preferred over *j*.

$$\Rightarrow$$
 M := { m_{ij} } $\in \mathbb{N}^{n \times n}$ and **B** := { b_{ij} } $\in \mathbb{N}^{n \times n}$.

Induced Probabilities by BTL Model

• We assume that the matrix $\mathbf{M} = \{m_{ij}\}$ is fixed a priori.

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Induced Probabilities by BTL Model

- We assume that the matrix $\mathbf{M} = \{m_{ij}\}$ is fixed a priori.
- The BTL model induces the following distributions:
 - 1 For fixed m_{ij} ,

$$p(b_{ij}|w_i,w_j) = \operatorname{Bin}(b_{ij};m_{ij},P_{ij}).$$

2 For fixed M,

$$p(\mathbf{B}|\boldsymbol{\lambda}) = \prod_{(i,j):i < j} \operatorname{Bin}(b_{ij}; m_{ij}, P_{ij}),$$

Bayesian BTL Model

■ Adopt the Bayesian BTL framework by Caron & Doucet⁵:

⁵F. Caron and A. Doucet, "Efficient Bayesian Inference for Generalized Bradley-Terry Models", in JCGS, 2012

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Bayesian BTL Model

■ Adopt the Bayesian BTL framework by Caron & Doucet⁵:

1 Prior distribution: They assign

$$p(w_i) = \operatorname{Gam}(w_i; \alpha_i, \beta_i)$$

to each item $i \in [n]$, where $\alpha = \{\alpha_i\}_{i=1}^n, \beta := \{\beta_i\}_{i=1}^n \in \mathbb{R}_{++}^n$.

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2 Latent random variables: They introduce $\mathbf{Z} := \{Z_{ij}\} \in \mathbb{R}^{n \times n}$

$$Z_{ij} = Z_{ji} := \sum_{s=1}^{m_{ij}} \min\{Y_{si}, Y_{sj}\},$$

for $i, j \in [n]$ such that i < j, where

 $Y_i \sim \operatorname{Exp}(w_i)$ & $Y_j \sim \operatorname{Exp}(w_j)$ such that $P_{ij} = \Pr[Y_i < Y_j]$.

Known as Thurstonian interpretation of the BTL model.

⁵F. Caron and A. Doucet, "Efficient Bayesian Inference for Generalized Bradley-Terry Models", in JCGS, 2012

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Induced Probabilities by Bayesian BTL Model

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2 Prior × Likelihood:

$$p(\mathbf{w}, \mathbf{B}) = p(\mathbf{w})p(\mathbf{B}|\mathbf{w}) = \prod_{i=1}^{n} \operatorname{Gam}(w_i; \alpha_i, \beta_i) \prod_{i < j} \operatorname{Bin}\left(b_{ij}; m_{ij}, \frac{w_i}{w_i + w_j}\right)$$

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$$p(Z_{ij}|w_i,w_j) = \operatorname{Gam}(Z_{ij};m_{ij},w_i+w_j).$$

4 Posterior:

$$p(\mathbf{w}|\mathbf{B}, \mathbf{Z}) = \prod_{i=1}^{n} \operatorname{Gam}(w_{i}; \alpha_{i} + b_{i}, \beta_{i} + Z_{i}).$$

where $b_{i} := \sum_{j \neq i} b_{ij}$ and $Z_{i} := \sum_{j \neq i} Z_{ij}.$

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For any $r \ge 1$, we define the family of Bayes risks for estimating w

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For any $r \ge 1$, we define the family of Bayes risks for estimating w

1 from only B as

$$R_{\mathrm{B}} := \inf_{\varphi} \mathbb{E} \Big[\|\mathbf{w} - \boldsymbol{\varphi}(\mathbf{B})\|_{r}^{r} \Big],$$

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$$R_{\mathrm{B}}^{*} := \inf_{\varphi^{*}} \mathbb{E}\bigg[\big\| \mathbf{w} - \varphi^{*}(\mathbf{B}, \mathbf{Z}) \big\|_{r}^{r} \bigg],$$

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$$R_{\rm B} \geq R_{\rm B}^*$$
.

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 $w_i \sim \text{Gam}(w_i; \alpha_i, \beta_i)$ Prior on w_i

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Image: Image:



$$P_{ij} = rac{w_i}{w_i + w_j}$$
 BTL model

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 $Y_{si} \sim \text{Exp}(w_i)$ Latent "Arrival Times"

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 $b_{ij} \sim \operatorname{Bin}(b_{ij}; m_{ij}, P_{ij})$ Num of times *i* beats *j* out of m_{ij} games

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$$Z_{ij} = \sum_{s=1}^{m_{ij}} \min\{Y_{si}, Y_{sj}\}:$$
 Latent variables

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Ranking

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 $arphi(\mathbf{B})$ and $arphi^*(\mathbf{B},\mathbf{Z})$: Functions to estimate w

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For r = 2, can compute the Bayesian Cramér-Rao bound on $R_{\rm B}$.

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 - **1** Theorem 3 of Xu and Raginsky⁶ reads: For any $r \ge 1$,

$$R_{\rm B}^* \geq \frac{n}{re} \left(V_n \cdot \Gamma \left(1 + \frac{n}{r} \right) \right)^{-r/n} \exp \left[-\frac{r}{n} \left(I(\mathbf{w}; \mathbf{B}, \mathbf{Z}) - h(\mathbf{w}) \right) \right],$$

where V_n is the volume of the unit ball in $(\mathbb{R}^n, \|\cdot\|_r)$.

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where V_n is the volume of the unit ball in $(\mathbb{R}^n, \|\cdot\|_r)$.

2 Using Stirling's approximation, we upper bound

$$I(\mathbf{w}; \mathbf{B}, \mathbf{Z}) - h(\mathbf{w}) = \mathbb{E}\left[\log p(\mathbf{w}|\mathbf{B}, \mathbf{Z})\right].$$

Family of Information-Theoretic Lower Bounds

Theorem

For all $i \in [n]$, let

$$m_i := rac{1}{2} \sum_{j \neq i} m_{ij}$$

half the total num. of games i plays

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Then, the Bayes risk is asymptotically lower bounded by

$$R_{\mathbf{B}} \gtrsim \frac{n}{re} \left(V_n \cdot \Gamma \left(1 + \frac{n}{r} \right) \right)^{-r/n} \exp \left[-r E(\mathbf{B}, \boldsymbol{\alpha}, \boldsymbol{\beta}) \right],$$

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where

$$E(\mathbf{B}, \boldsymbol{\alpha}, \boldsymbol{\beta}) := \sum_{i=1}^{n} \left(-\frac{1}{2} \log (2\pi) + \log \beta_i - \psi(\alpha_i) + \frac{1}{2} \log (\alpha_i + m_i) \right).$$

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Information-Theoretic Lower Bounds

Take $\alpha_i = \alpha$ and $\beta_i = \beta$, for each $i \in [n]$.

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Information-Theoretic Lower Bounds

- Take $\alpha_i = \alpha$ and $\beta_i = \beta$, for each $i \in [n]$.
- For the L^1 norm (r = 1),

$$R_{\rm B}^* \gtrsim \sqrt{\frac{\pi}{2}} \exp\left[-\left(\log \beta - \psi(\alpha) + 1\right)\right] \frac{n}{\sqrt{\alpha/n+m}},$$

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• For the squared L^2 norm (r = 2),

$$R_{\rm B} \gtrsim \exp\left[-2(\log\beta - \psi(\alpha)) - 1\right] \frac{n}{\alpha/n+m}.$$

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Performance of Lower Bounds: L^1 error



Figure: L^1 error of the EM algo. and the information-theoretic lower bound (for n = 100, $\alpha = 5$ and $\beta = \alpha n - 1$).

Perf. of Lower Bounds: MSE (squared L^2 error)



Figure: L^2 error of the EM algo., the IT lower bound and the BCRB (for $n = 100, \alpha = 5$, and $\beta = \alpha n - 1$).

Given a fixed budget of $m = \sum_{i \neq j} m_{ij}$ games,

how to allocate games among *n* players to minimize the bounds?

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Corollary (Optimal Connected Graphs)

Regular Connected Graphs are Optimal!



Proof:

Minimizing the lower bound is equivalent to maximizing

$$f({m_i}_{i \in [n]}) := \sum_{i=1}^n \frac{1}{2} \log (\alpha_i + m_i)$$

subject to $\sum_{i=1}^{n} m_i = m$ and $m_i \in \mathbb{N}$.

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Solution given by water-filling formula:

$$m_i = |\mu - \alpha_i|_+, \quad \forall i \in [n],$$

where $\mu > 0$ is chosen such that

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But when $\alpha_i = \alpha$ for all *i*, *m_i* are all equal.

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The Gamma Distribution with Fixed $\beta = 1$

$\alpha_i \uparrow \Longrightarrow$ Greater belief that $w_i \uparrow$ \implies Games *i* plays with others $m_i \downarrow$



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Corollary (Optimal Tree Graphs)



Corollary (Optimal Tree Graphs)



- 1 Best: Minimizes the (lower bound on the) Bayes Risk
- 2 Worst: Maximizes the (lower bound on the) Bayes Risk

Proof for Star:

Maximizing the lower bound on Bayes risk equivalent to

$$\min_{\mathbf{m}:\sum_{i}m_{i}=m}g(\mathbf{m}):=\frac{1}{2}\log\left(\alpha+2m+\sum_{i'\neq i^{*}}m_{i'}\right)+\sum_{i}\frac{1}{2}\log(\alpha+m_{i})$$

where $\mathbf{m} = \{m_i\}_{i \in [n]}$ and $i^* = 1$ is the central node.

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where $\mathbf{m} = \{m_i\}_{i \in [n]}$ and $i^* = 1$ is the central node.

Shift part of weight of an edge $m_{1j} > 0$, for $j \neq 1$, to create a new edge with weight m_{ji} such that $i \neq 1$. Can show that

$$\frac{\partial g(m_1,\ldots,m_n)}{\partial m_i} > 0$$

implying that f will be increased by the new configuration.

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Effect of Tree Graph Structure on IT Bound



Figure: IT bounds for diff. graph structures (for n = 100, $\alpha = 5$, $\beta = \alpha n - 1$).
Effect of Tree Graph Structure on BCRB



Figure: BCRB for diff. graph structures (for n = 100, $\alpha = 5$, $\beta = \alpha n - 1$).

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Final Remarks

Also derived lower bounds for the home-field advantage scenario:

$$P_{ij} = \begin{cases} Q_{ij} := \frac{\theta \lambda_i}{\theta \lambda_i + \lambda_j}, & \text{if } i \text{ is home}, \\ \overline{Q}_{ij} := \frac{\lambda_i}{\lambda_i + \theta \lambda_j}, & \text{if } j \text{ is home}, \end{cases}$$

where $\theta \in \mathbb{R}_{++}$ models the strength of advantage ($\theta > 1$)

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Future works: Matching information-theoretic upper bounds

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- Other questions related to comparing graph structure, e.g.,

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