

Adversarial top- K ranking

Vincent Tan
ECE and Mathematics, NUS

2016 Shannon Workshop, SJTU
December 15, 2016



Changho Suh @KAIST



Renbo Zhao @NUS

Ranking

One of the fundamental problems in a wide range of contexts.

Applications: web search, recommendation systems,
social choice, sports competitions, voting, etc

Lots of efforts made in developing ranking algorithms.

A variety of statistical **models** introduced for evaluating ranking algorithms.

Ranking : An Example and Difficulties

Example: web search

Google YAHOO! bing



$n = 10^9$ websites

$$\binom{n}{2} \approx n^2 = 10^{18} \text{ comparisons}$$

Really need $\geq n^2$ comparisons?

Large Scale Ranking

Suppose (i) we want a **total ordering** &
(ii) pairwise comparisons are **randomly given**
(probabilistically).

→ Requires $\geq n^2$ comparisons.

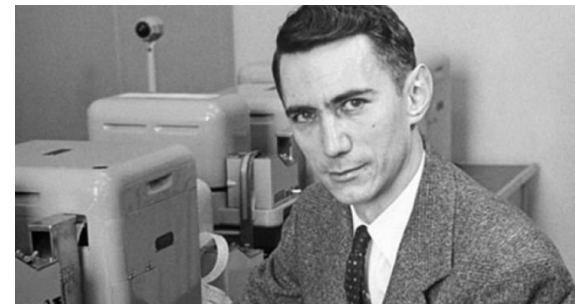


No way to identify the ordering between 1 and 2
without the direct comparison.

Comparison must be made **with probability 1**.

Things are even worse if one has **noisy** data.

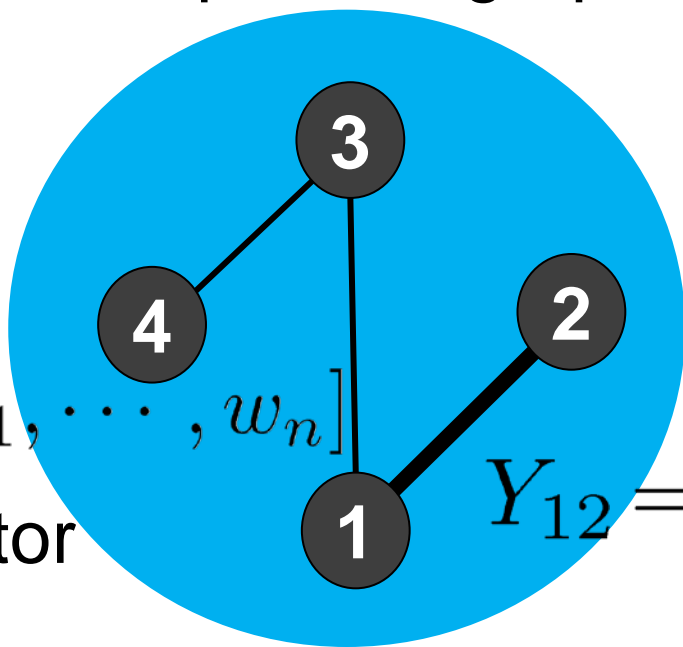
Solution: Shannon-theoretic approach



A prominent model

Bradley-Terry-Luce '52

comparison graph

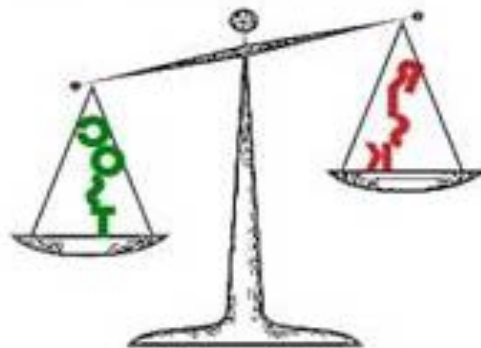


$\mathbf{w} = [w_1, \dots, w_n]$
score vector

$$Y_{12} = \mathbb{I}\{\text{item 1} \succ \text{item 2}\}$$

$$\sim \text{Bern} \left(\frac{w_1}{w_1 + w_2} \right)$$

L independent copies: $Y_{12}^{(1)}, \dots, Y_{12}^{(L)}$



Challenge in crowdsourced settings



Spammers provide answers in an **adversarial** manner.
The BTL does not respect the **adversarial** scenario.

Adversarial BTL model

Adapted from Chen et al. '13

faithful
population
 η
portion



$$Y_{ij} \sim \text{Bern} \left(\frac{w_i}{w_i + w_j} \right) \quad \text{w.p. } \eta$$

adversarial
population
 $1 - \eta$
portion



$$Y_{ij} \sim \text{Bern} \left(\frac{w_j/w_i}{w_j/w_i + 1} \right) \quad \text{w.p. } 1 - \eta$$

$$Y_{ij} \sim \text{Bern} \left(\eta \cdot \frac{w_i}{w_i + w_j} + (1 - \eta) \cdot \frac{w_j}{w_i + w_j} \right)$$

WLOG assume $\eta \in (0.5, 1]$

Related work: Crowdsourced BTL

[Chen et al. '13]

Given an observed pair, each sample has **different** distributions.

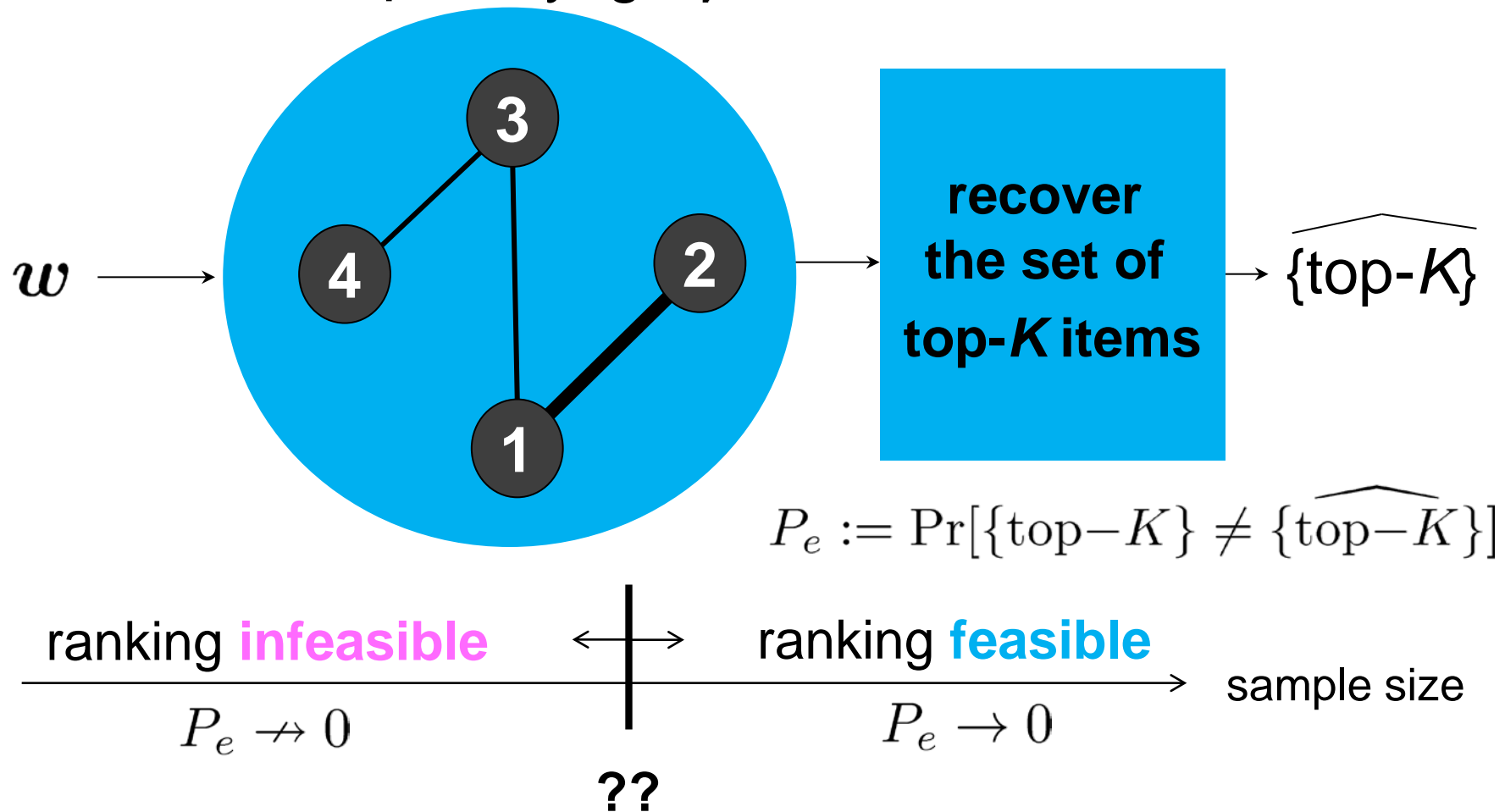
$$Y_{ij}^{(\ell)} \sim \text{Bern} \left(\underset{\substack{\nearrow \\ \text{quality parameter}}}{\eta_\ell} \cdot \frac{w_i}{w_i + w_j} + (1 - \eta_\ell) \cdot \frac{w_j}{w_i + w_j} \right)$$

Subsumes as a special case our adversarial BTL model.

Developed a ranking algorithm **but without theoretical guarantees.**

Top-K ranking

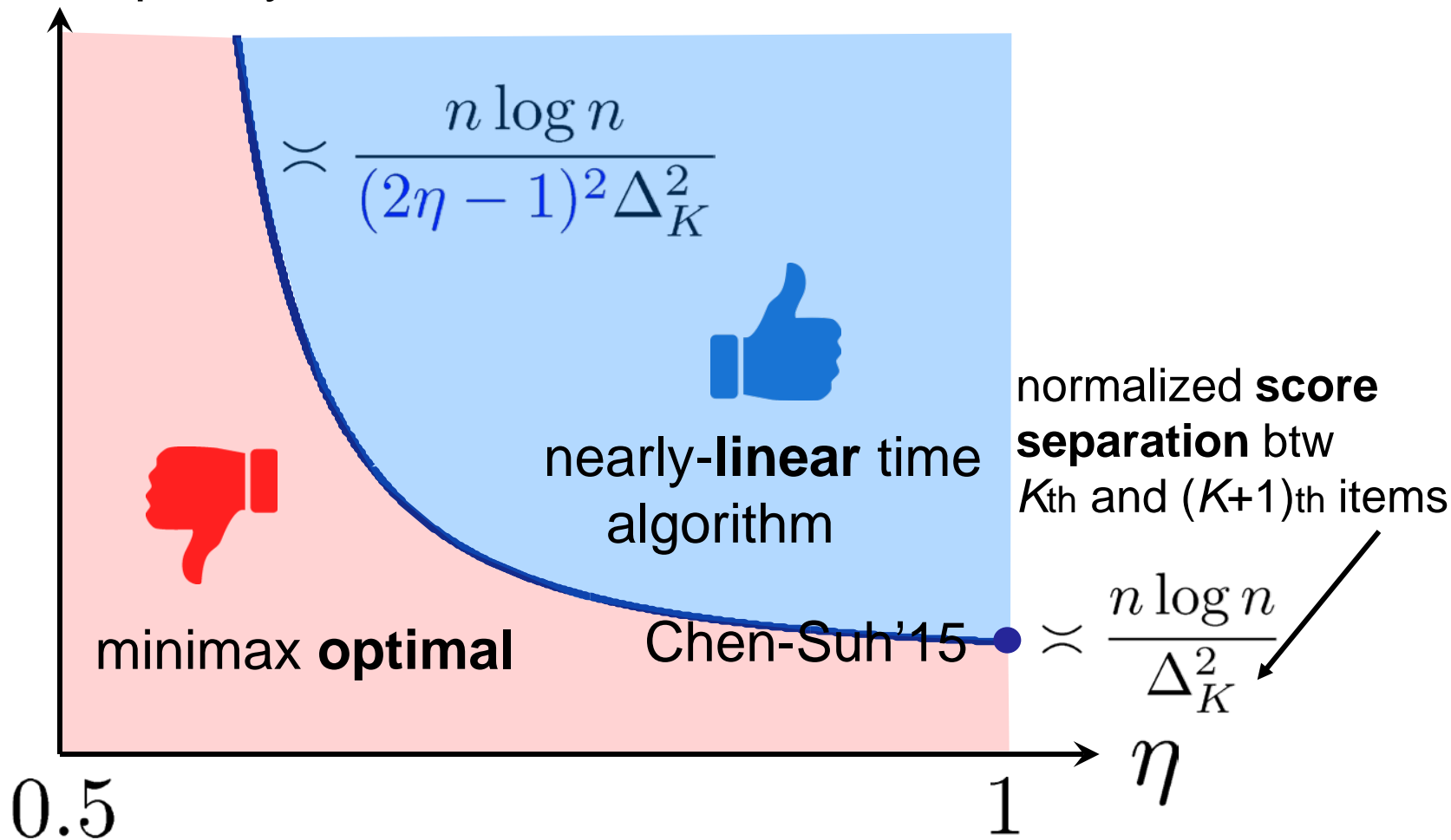
Erkko-Petri graph



Contribution #1

η -known

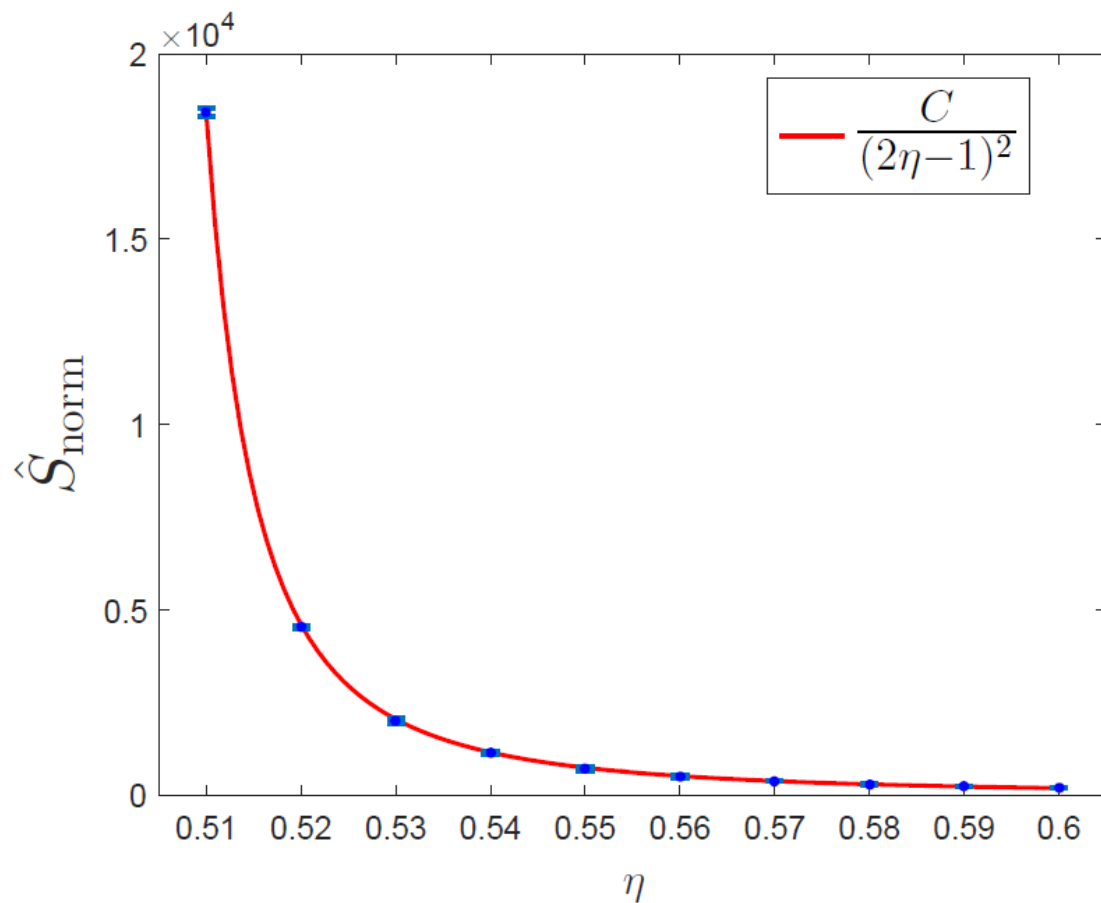
sample complexity



Contribution #1

η -known

Experimental Result



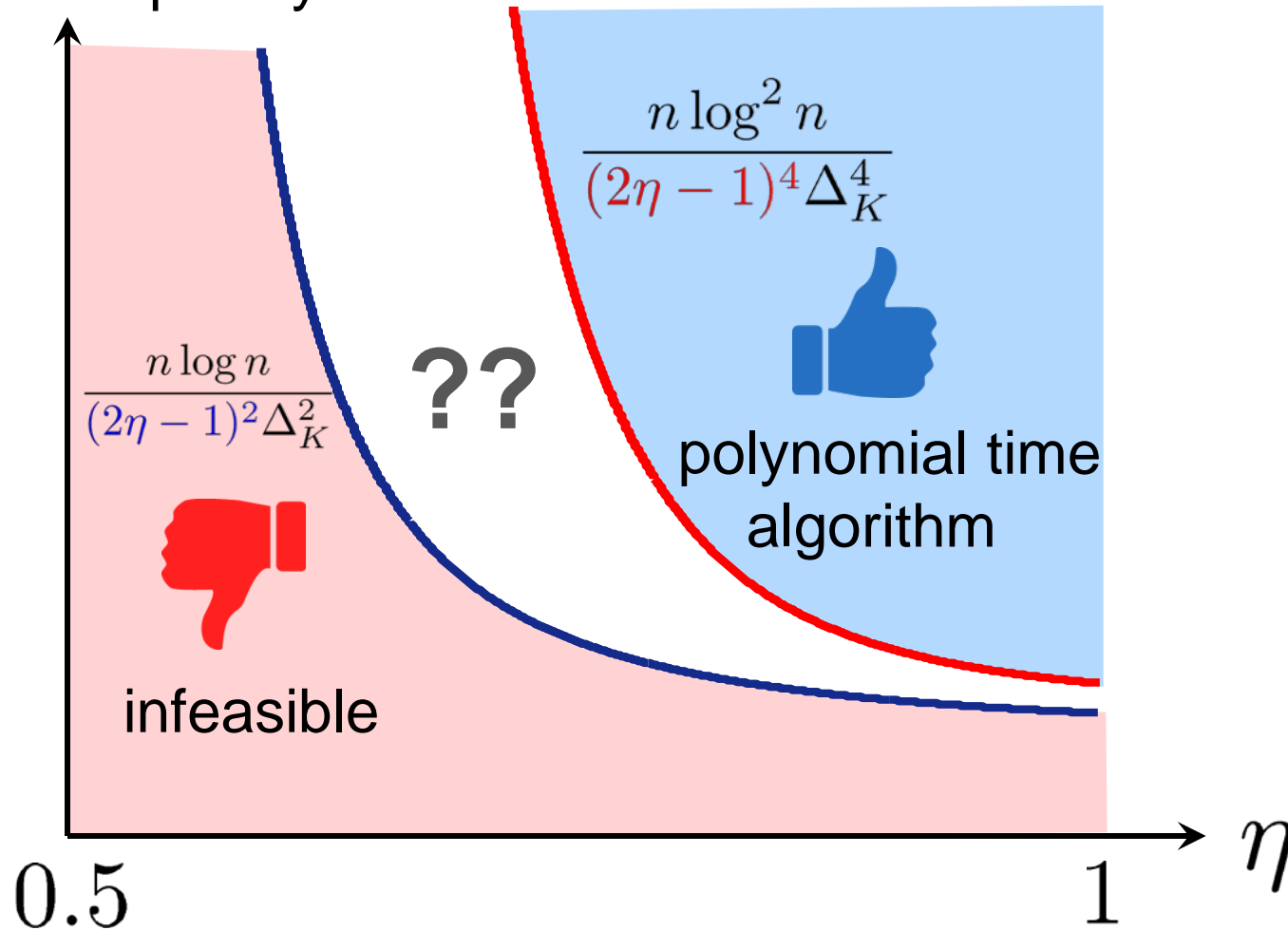
$$\hat{S} := \binom{n}{2} p \hat{L}_{\text{ave}}$$

$$\hat{S}_{\text{norm}} := \frac{\hat{S}}{(n \log n) / \Delta_K^2}$$

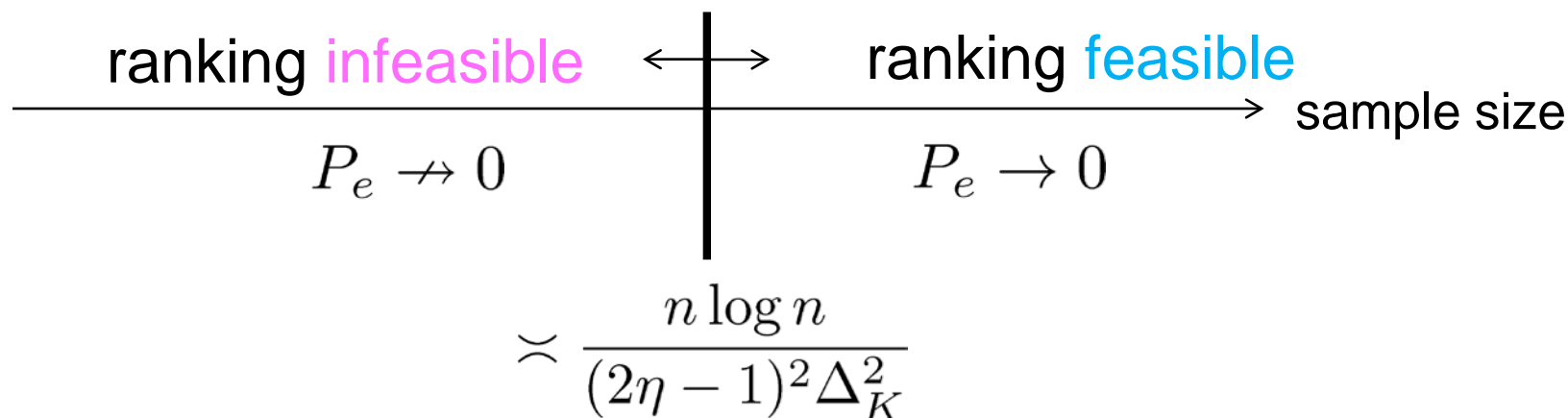
Contribution #2

η -unknown

sample complexity



Optimality

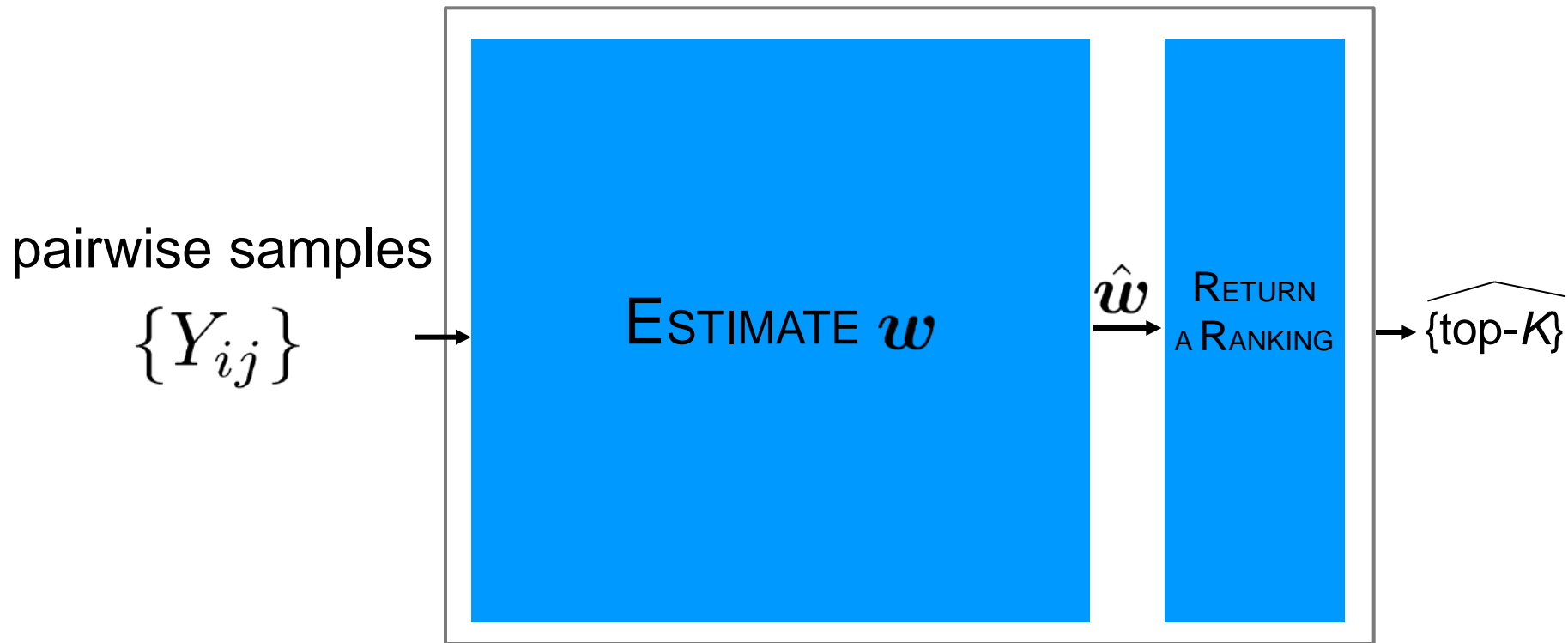


- **Minimax** optimality:
Construction of the worst-case score vector
- Translation to **hypothesis testing**:
Construction of multiple hypotheses

Tools: KL divergence, Generalized Fano's inequality,
Reverse Pinsker inequality

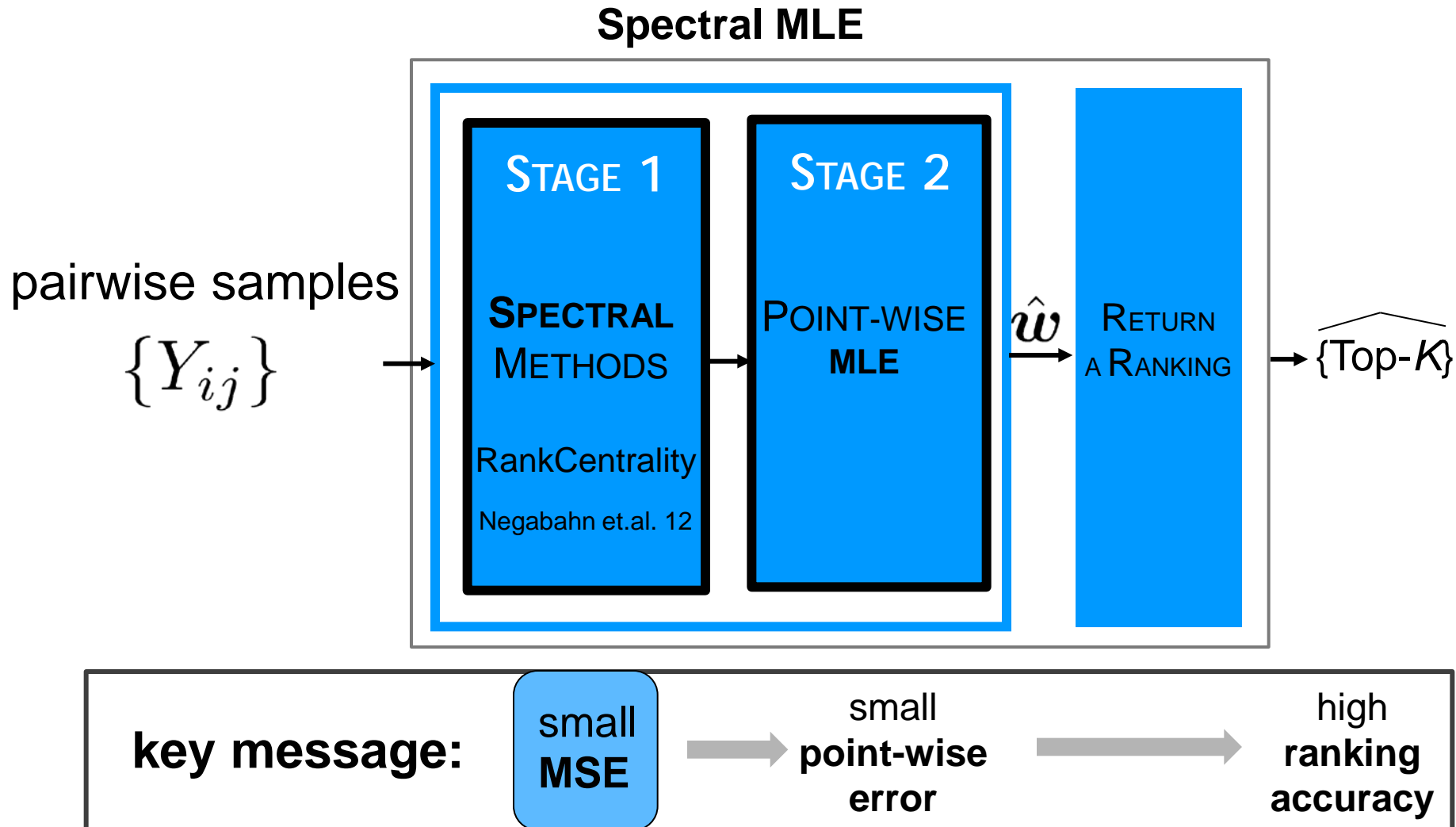
Ranking algorithm

η -known



Remember: Scores determine a ranking.

→ Take a **two-step** approach.



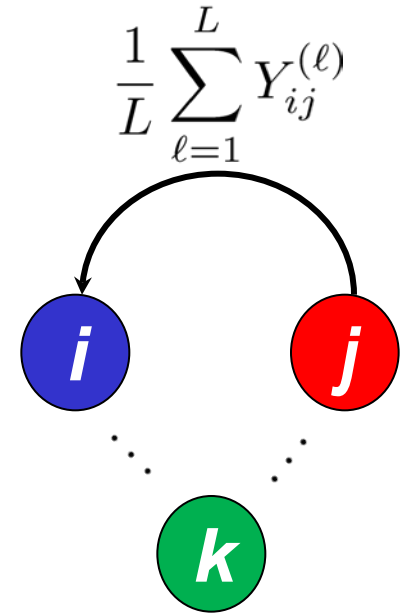
How to ensure small MSE?

$$\eta \in (0.5, 1]$$

Recall $\eta = 1$:

L ind. copies $Y_{ij}^{(1)}, Y_{ij}^{(2)}, \dots, Y_{ij}^{(L)}$

$$\frac{1}{L} \sum_{\ell=1}^L Y_{ij}^{(\ell)} \longrightarrow \frac{w_i}{w_i + w_j}$$



stationary dist. \rightarrow \mathbf{w} (up to const. scaling)

How to ensure small MSE?

$$\eta \in (0.5, 1]$$

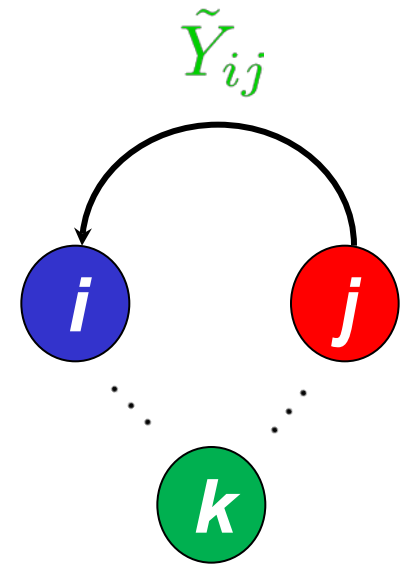
Arbitrary η

L ind. copies $Y_{ij}^{(1)}, Y_{ij}^{(2)}, \dots, Y_{ij}^{(L)}$

$$\begin{aligned} \frac{1}{L} \sum_{\ell=1}^L Y_{ij}^{(\ell)} &\longrightarrow \eta \cdot \frac{w_i}{w_i + w_j} + (1 - \eta) \cdot \frac{w_j}{w_i + w_j} \\ &= (2\eta - 1) \cdot \frac{w_i}{w_i + w_j} + 1 - \eta \end{aligned}$$

$\tilde{Y}_{ij} :=$

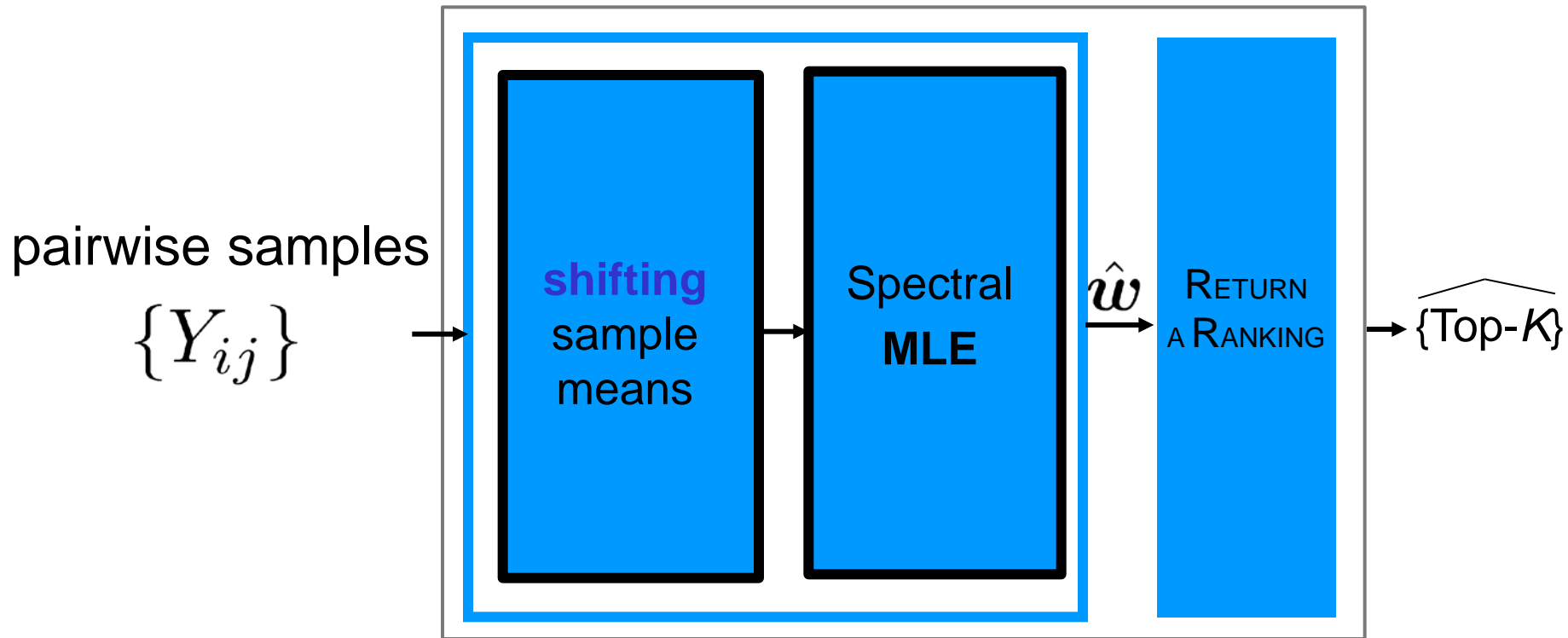
$$\boxed{\frac{\frac{1}{L} \sum_{\ell=1}^L Y_{ij}^{(\ell)} - (1 - \eta)}{2\eta - 1}} \longrightarrow \frac{w_i}{w_i + w_j}$$



Idea: Construct Markov Chain now with $\{\tilde{Y}_{ij}\}$

Ranking algorithm

η -known $\eta \in (0.5, 1]$



Using several ineq. (Hoeffding, Bernstein, Tropp ...), showed:

$$\text{sample size} \gtrsim \frac{n \log n}{(2\eta - 1)^2 \Delta_K^2} \longrightarrow \text{feasible ranking}$$

What if η is **unknown**?

Adversarial BTL model is a **mixture model**.

Mixture model learning problems are difficult in general.

Recent development:

Tensor methods: Jain-Oh '14, Anandkumar et al. '14

Key insight: exact 2nd & 3rd moments \rightarrow sufficient statistics

Our setting:

Can obtain **estimates** of 2nd & 3rd moments

\rightarrow Can estimate η

Unknown η ? High-level Algorithm

1) Turn weights into **distribution vectors**

$$\pi_0 := \left[\cdots \quad \frac{w_i}{w_i + w_j} \quad \frac{w_j}{w_i + w_j} \quad \frac{w_{i'}}{w_{i'} + w_{j'}} \quad \frac{w_{j'}}{w_{i'} + w_{j'}} \quad \cdots \right]^T$$

2) Estimate **moments**

$$M_2 := \eta \pi_0 \otimes \pi_0 + (1 - \eta) \pi_1 \otimes \pi_1,$$

$$M_3 := \eta \pi_0 \otimes \pi_0 \otimes \pi_0 + (1 - \eta) \pi_1 \otimes \pi_1 \otimes \pi_1,$$

3) Solve a **Least Squares Problem**

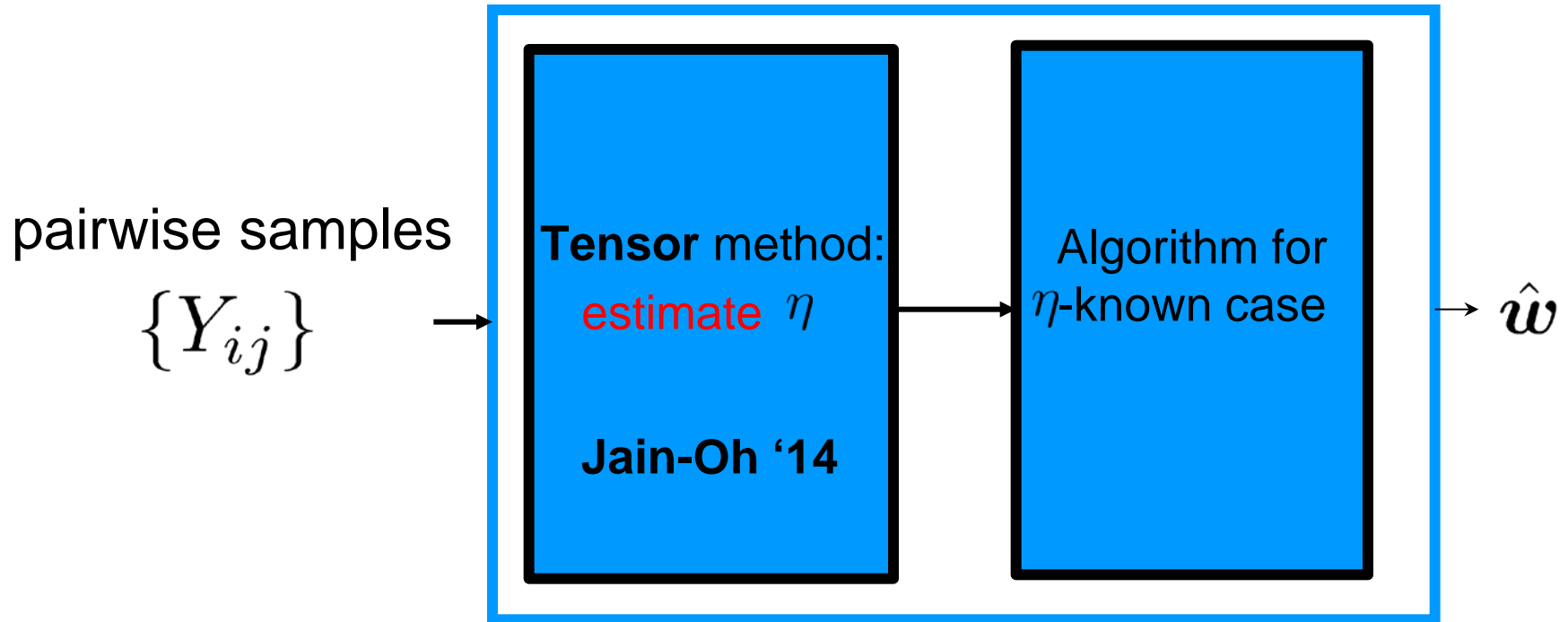
$$\hat{G} \in \arg \min_{Z \in \mathbb{R}^{2 \times 2 \times 2}} \left\| \mathcal{P}_{\Omega_3} \left(Z \left[P_{\hat{M}_2} \right]_3 - \frac{1}{|\mathcal{I}_2|} \sum_{t \in \mathcal{I}_2} \otimes^3 \underline{Y}^{(t)} \right) \left[Q_{\hat{M}_2} \right]_3 \right\|_F^2.$$

4) Find **leading eigenvalue** λ_1 of \hat{G} which is related to the mixing weight as follows:

$$\hat{\eta} = \lambda_1^{-2}$$

Ranking algorithm

η -unknown



How does the η -estimation affect sample complexity?

Tradeoff btw $|\eta - \hat{\eta}|$ & sample complexity

With very careful analysis, we can derive a lemma:

$$|\eta - \hat{\eta}| \leq \epsilon \longrightarrow \text{sample size required} \gtrsim \frac{n \log^2 n}{\epsilon^2}$$

$$|\eta - \hat{\eta}| \downarrow \longrightarrow \|\mathbf{w} - \hat{\mathbf{w}}\|_{\infty} \downarrow \text{ but sample size required } \uparrow$$

$$|\eta - \hat{\eta}| \uparrow \longrightarrow \text{sample size required } \downarrow \text{ but } \|\mathbf{w} - \hat{\mathbf{w}}\|_{\infty} \uparrow$$

We could find a **sweet spot** to show that

$$\text{sample size} \gtrsim \frac{n \log^2 n}{(2\eta - 1)^4 \Delta_K^4} \longrightarrow \text{feasible ranking}$$

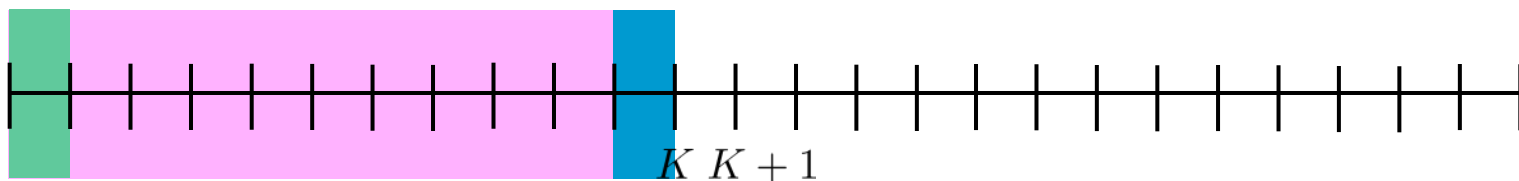
Conclusion

- Explored a **top- K** ranking problem for an **adversarial** setting.
- Characterized order-wise optimal **sample complexity** for η -**known** case.
- Established an **upper bound** on the sample complexity for η -**unknown** case.
- Developed computationally **efficient** algorithms.
- Full version of paper on arXiv 1602.04567.

backup slides

Extension

- Detailed ranking among the top- K items



$$\text{sample size required} \asymp \frac{n \log n}{\Delta_K^2}$$

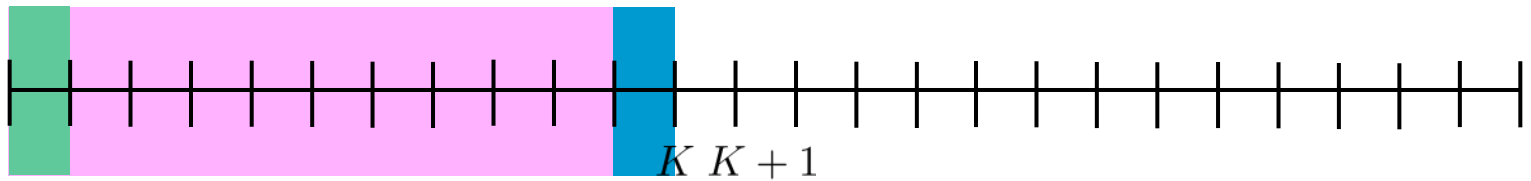
$$\text{sample size required} \asymp \frac{n \log n}{\Delta_{K-1}^2}$$

\vdots

$$\text{sample size required} \asymp \frac{n \log n}{\Delta_1^2}$$

Extension

- Detailed ranking among the top- K items



$$\text{sample size required} \gtrsim \frac{n \log n}{\min \Delta_i^2}$$

- Can easily extend to any- K ranking.

Related work: General comp. model

No ground truth score vector

[Shah-Wainwright`15]

Instead we are given: $p_{ij} := \Pr [\text{item } i \succ \text{item } j]$

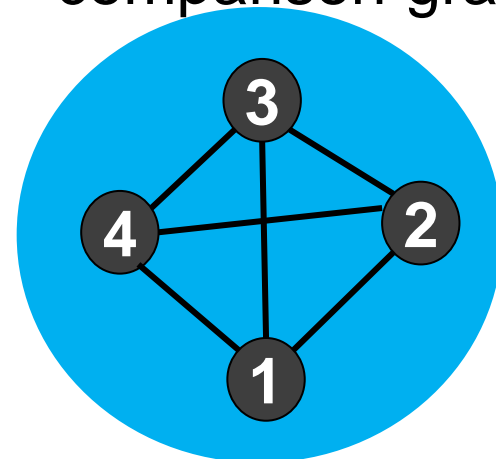
Subsumes as special cases many parametric models
(including BTL and ours)

Assumption:

of comparisons per every edge

$$\sim \text{Binom}(r, p)$$

comparison graph



Random vs. adaptive sampling

Case 1: $\Delta_1 \approx \Delta_2 \approx \dots \approx \Delta_{n-1}$

Both sampling methods yield almost the same performance.

Case 2: $\Delta_1 \ll \Delta_i \ (i \geq 2)$

Adaptive sampling outperforms random sampling.